

# Competition Avoidance vs Herding in Job Search: Evidence from Large-scale Field Experiments on an Online Job Board

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We study how information that may simultaneously signal the degree of competition and vacancy quality affects job search. To do so, we conduct three experiments on a large online job platform in which the treatment varies what information is shown to job seekers. Information about the number of prior applicants to a vacancy increases the number of applications and redirects them to vacancies with few prior applications. Information about vacancy age increases application rates, especially to new vacancies. To further investigate the causal mechanisms, we conduct and analyze a survey choice experiment. We conclude that job seekers prefer to avoid competition rather than using the popularity of a vacancy as a signal of quality.

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## 1. Introduction

People use information about the behavior of others to make decisions about what to consume, where to apply, and how to search. For example, signals of song popularity (downloads) caused participants to listen to more popular songs ([Salganik et al. \(2006\)](#)). Herding behavior influenced by prior popularity signals has also been observed in social media, microlending, crowdfunding, petitions, job acceptance, and online platform adoption.<sup>1</sup> People may follow signals of popularity for a variety of reasons, including because these signals contain information about the quality of an option. However, in labor, housing, and other constrained markets, popularity is also correlated with the degree of competition. More competitive options have a lower likelihood of success because the number of slots is fixed and most vacancies are filled from early applicants ([Van Ours and Ridder \(1992\)](#)). This creates a tradeoff for individuals - apply to a more popular option but have a lower chance of success.

Platform designers must choose whether and how to display signals of popularity and competition information to its users, and this decision critically depends on how users perceive that

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<sup>1</sup> See [Coffman et al. \(2017\)](#), [Muchnik et al. \(2013\)](#), [Tucker and Zhang \(2010\)](#), [Van de Rijt et al. \(2014\)](#), and [Zhang and Liu \(2012\)](#).

information. The provision of information may improve market efficiency if it causes searchers to redirect applications to less competitive options, as predicted by some models of search (Wright et al. 2019). It may also improve market efficiency if it directs users to higher quality options. On the other hand, popularity information may be harmful to users if it induces herding and results in wasted search. Which of these effects dominates is unknown, and major job search platforms differ in their information designs.<sup>2</sup>

We investigate how job seekers use information that simultaneously conveys vacancy popularity and competition in the labor market. We find that, when given information of the number of prior applicants to a vacancy, job seekers redirect their search towards vacancies with few prior applications. In a related treatment, we find that job seekers prefer recently posted jobs to jobs that are on the platform for a longer duration. Our main results demonstrate that participants in the labor market prefer to avoid competition even if it means applying to less popular jobs.

To conduct our study, we use three experiments encompassing millions of job seekers conducted on an online job board. We find that information about the vacancy level degree of competition affects job search, and in particular, directs applications towards vacancies with few applications in line with the spirit of directed search models. The applications induced by the information do not fully crowd out other applications and exhibit similar outcomes to those in the control group.

The setting of our study is a job board operated by Facebook called ‘Jobs on Facebook’ (JOF) until 2022.<sup>3</sup> JOF was global, and mainly catered to full-time positions that did not require a college education. Users of the platform saw a list of vacancies and had access to a rich set of filters by which to refine their search. Searchers had the ability to apply to most vacancies using the platform. Consequently, the platform was also able to directly observe the number of applications sent to a particular vacancy in real-time and to display information about the count of prior applications to searchers.

We begin the paper by proposing a simple theoretical framework for a job seeker’s choice of whether to apply for a vacancy or not. Information can affect two components of the seeker’s choice, beliefs about job quality and beliefs about the likelihood of getting the job conditional on applying. Whether a particular piece of information increases or decreases application rates depends on how it affects these two components and is the empirical question we try to answer with our experiments.

<sup>2</sup> As of 2022, Indeed and Google Jobs do not display the number of other applications on the search page while LinkedIn does. AngelList Jobs does not display the number of prior applicants but does display that the employer is actively hiring. All platforms display vacancy age, but EconJobMarket also displays the deadline for applications.

<sup>3</sup> In December of 2022 Meta made an announcement about phasing out JOF. See here: <https://www.facebook.com/business/help/982945655901961>.

We also document two motivating facts. First, we show that being an earlier applicant increased the likelihood that an application was viewed, responded to, and interviewed by an employer. Second, we show that applications to recently posted vacancies experienced higher employer response rates. This relationship between employer response and vacancy age holds true even conditional on application order. These facts point to a potential role for information signals in directing job search.

We then discuss our experimental evidence regarding the effects of competition signals. Starting in March 2019 and continuing through August of 2019, Facebook conducted three experiments related to our research questions. All three experiments contained treatment arms that displayed information about the number of prior applicants to a vacancy in the search interface. The different treatment arms and experiments varied the frequency (every vacancy, every three vacancies, or every 10 vacancies) and color (grey vs blue) of the information.

We find that these treatments increased application rates to vacancies with fewer than five prior applications by 3.8%, with a range of .9% to 6.4% depending on the experiment. In contrast, application rates to vacancies with many prior applications fell. We also find that the total number of application increased due to the treatment. Therefore, the additional applications to low competition vacancies did not fully crowd out existing applications. The frequency and the color of information did not have first order effects on the rate of applications to these vacancies.

To better understand the mechanisms behind our finding, we conducted a choice survey in which we ask online panel respondents to make choices over vacancies that vary in their wages and number of prior applications. We find that many respondents preferred to apply to vacancies that have fewer applications even when the vacancy has a low wage. In contrast, very few respondents preferred a job with a low wage when the number of prior applicants is high. The survey analysis corroborates our field experimental finding that job seekers' responses to competition information are driven by competition avoidance and not by herding.

Next, we consider the role of vacancy age in directing job searchers towards less congested vacancies. Since vacancy age is positively correlated with the number of applications, it is potentially a proxy for congestion. Furthermore, since vacancy age information is available to job-seekers in many online settings, it is potentially a driver of directed search. The default JOF interface displayed vacancy age, but we removed this information in one of our experiments. In contrast to our other treatments in which the platform added signals to help direct search, removing vacancy age potentially removed a signal that otherwise could be used to direct search. But removing vacancy age helps us better understand what job-seekers were trying to accomplish.

Workers used vacancy age to decide whether and where to apply. Job seekers who did not have information about vacancy age clicked on 3% fewer vacancies and sent 1.8% fewer applications.

The removal of vacancy age also had distributional effects. Treated users were less likely to apply to new vacancies and were more likely to apply to old vacancies. We also find that removing vacancy age increased the concentration of applications.

Lastly, we consider the causal effect of the treatment on the success and quality of applications submitted. On the one hand, applications in the treatment group should benefit from being earlier. On the other hand, searchers may choose to send these applications to worse matching vacancies which would then have a lower likelihood of resulting in a hire. As a result, the sign of effect of the treatment on application outcomes is theoretically ambiguous. We find that applications in the treatment were not substantially more likely to be viewed, contacted, or hired. This suggests that treated applications that any negative selection of applications in the treatment was not large. Furthermore, since the total number of applications increased but the the outcomes were not harmed, the treatment likely helped treated applicants.

We do not find that behavior in the job market exhibits the type of social contagion based on popularity signals found in other digital settings ([Salganik et al. \(2006\)](#), [Muchnik et al. \(2013\)](#)). In other words, there is no danger of indicating a job is “popular” causing it to receive even more applications, as in some kind of social learning or information cascade scenario. In short, job-seekers view the application process more as a congestion game and all else equal, would prefer facing fewer competitors.

Our results provide causal evidence about platform generated signals that can direct search. The competition avoidance due to the signals corresponds to a common feature of directed search models, that workers care about the likelihood that their application is successful ([Wright et al. \(2019\)](#)). [Cheron and Decreuse \(2017\)](#) and [Albrecht et al. \(2017\)](#) focus specifically on the importance of ‘phantom vacancies’, which are vacancies that have already been filled. In their models, workers rationally direct search towards newer postings—a phenomenon which we confirm using experimental variation regarding information about vacancy age.

Other papers in the literature have shown that search is directed on many vacancy characteristics such as compensation ([Belot et al. \(2018\)](#), [Banfi and Villena-Roldan \(2019\)](#), [Flory et al. \(2015\)](#), [Samek \(2019\)](#)) and signals of employer preferences ([Kuhn et al. \(2020\)](#), [Leibbrandt and List \(2018\)](#), [Ibañez and Riener \(2018\)](#)). A particular focus of this literature has been gender differences in preferences regarding the competitiveness of compensation schemes. We study a different aspect of competitiveness, which is the amount of competition to get hired. We are not able to detect differences between men and women regarding their preferences towards such vacancies.

Our treatments are enabled by the fact that digital job boards have a bird’s eye view of the market. While this may seemingly limit the applicability of this approach, given that the labor market as a whole is decentralized, an increasing amount of job search occurred on digital job boards

(Kuhn and Mansour 2014, Baker and Fradkin 2017, Kroft and Pope 2014, Marinescu 2017). These job boards make decisions that could ameliorate—or worsen—congestion. Due to the heterogeneity in preferences for vacancies-across seekers and vacancies, centralized matching is infeasible. Instead, the platform can influence matching indirectly through the information it chooses to display and emphasize.

We build on the paper by Gee (2019), who varied whether the number of ‘people who clicked to apply’ to a vacancy was shown on LinkedIn. A critical difference between our work and Gee’s is that our experiment is conducted at a different point in the job application “funnel.” In our setting, job-seekers can scan over a collection of jobs and learn about the relative degree of competition, whereas in Gee (2019) they only learn about competition after choosing to investigate a particular job.

The timing of information acquisition has previously been shown to matter greatly for outcomes in search markets (Branco et al. (2012), Hodgson and Lewis (2020), Gardete and Hunter (2020), Abaluck and Compiani (2020)). This is the likely reason that we find that applications increased for vacancies with few prior applicants and decreased for vacancies with many applicants while Gee (2019) does not. When comparing other outcomes, Gee (2019) finds that applications per view increase and overall applications increase by 3.7%. In comparison, we find that applications increase by less than 1% but that applications per detailed view increase by a similar amount as in Gee (2019). Although the timing of information acquisition is a key detail, there are also other differences between our paper and that of Gee (2019) that may contribute to any differences in findings. In particular, our sample tends to be less educated and less likely to be non-US based (42% US based in Gee (2019) vs 20% to 27% in our experiments).

Several other papers have used data on search in digital labor platforms. Faberman and Kudlyak (2019) and Marinescu (2017) study how search evolves over the course of an unemployment spell, Marinescu (2017), Marinescu and Skandalis (2021), and Baker and Fradkin (2017) study how unemployment insurance affects online job search. Azar et al. (2019) use data from CareerBuilder to build a demand model of applications and use it to estimate firms’ market power in the labor market. Le Barbanchon et al. (2021) use data on search criteria declared to a public employment agency, as well as applications on a digital platform, to show that women care more than men about commuting when it comes to applying for jobs. Skandalis (2018) shows how job search is affected by news about a company’s hiring needs.

The above literature has not focused on the market design possibilities available on job boards. Our paper shows that the information design of these job boards has large effects on behavior. This creates opportunities for additional information interventions, and the study of their equilibrium effects using market level experiments. In this way, the market design innovations pioneered in

other digital platforms, such as those for labor procurement (Horton (2017)), dating (Fong (2019)) and accommodations (Fradkin (2017)), can be used to improve outcomes on digital job boards. More recently, Hensvik et al. (2020) have studied the effect of algorithmic recommendations in a job board similar to ours.

The rest of the paper is organized as follows. Section 2 describes the empirical context. Section 3 presents a decision framework and stylized facts that motivate our experimental treatments. Section 4 discusses the design of our various interventions. Section 5 reports the effects of the treatment on search behavior. Section 6 discusses our results with respect to vacancy age. Section 7 reports on the differences between applications across treatments. Section 8 concludes.

## 2. Empirical context

JOF was an online job board that operated between 2016 and 2022. The product was global in nature and mainly catered to positions which do not require a college education. The share of US users across our experiments ranged from 21% to 25%, and the median user in our experiments was between 31 and 33 years old across experiments.<sup>4</sup> Employers posted vacancies and job-seekers browsed vacancies and sent applications through the platform. The service was free for both sides, but job-seekers needed a Facebook account. A posted vacancy was automatically live for 30 days, but employers could renew it manually. This resulted in an average duration during our testing period of 42 days.<sup>5</sup> Even before the launch of JOF, there was substantial job-search behavior on Facebook (Gee et al. 2017).

Job-seekers were exposed to JOF via the “News Feed” and via notifications.<sup>6</sup> They could also navigate to JOF by clicking on the “Explore” tab and then clicking on a briefcase icon labeled “Jobs.” The JOF interface was similar to other job boards, though most of job search occurred on mobile devices. That most use occurred on mobile presents opportunities—if the user had enabled location-tracking, vacancies within a given radius could easily be shown—but also challenges, in that there was a constrained space in which to present information.

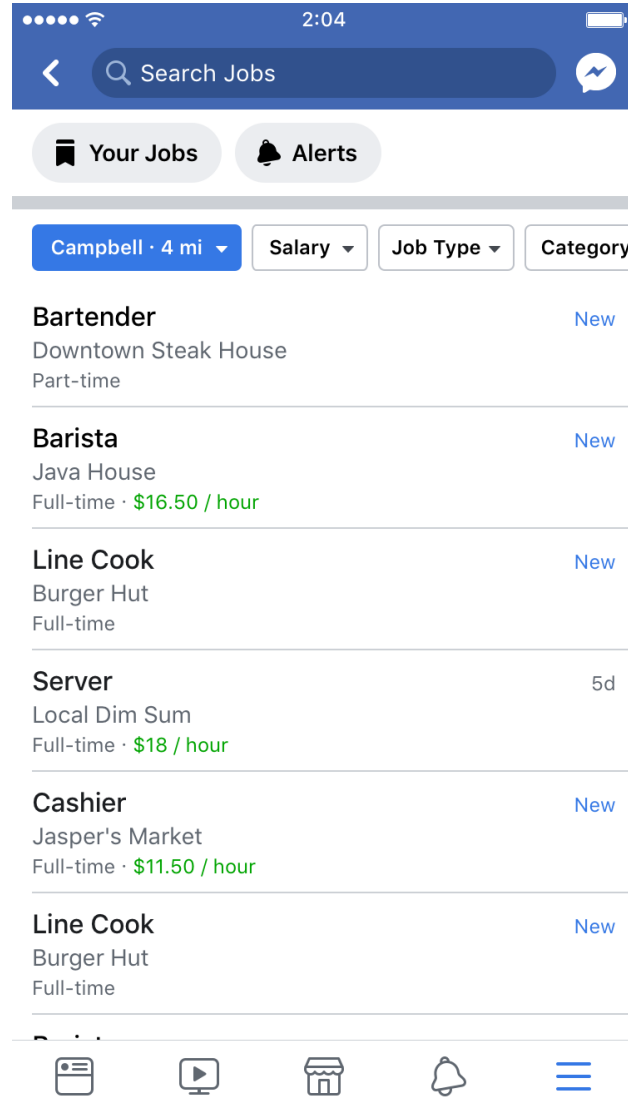
When looking for work, job-seekers could enter a number of criteria to narrow their search, including their location and the type of position they are interested in. Figure 1 shows the *status quo* job search interface (the job board)—as we will discuss at length, this presentation was modified

<sup>4</sup> According to publicly available data, approximately 10% of Facebook’s overall user base is based in the United States. JOF’s relatively high penetration in the US is likely due to the fact that JOF first launched in the US prior to being expanded globally (<https://perma.cc/6CAM-EKLR>).

<sup>5</sup> We calculate this duration for all vacancies created between March 2019 and August 2019. Note that Facebook reviewed job postings and may remove them from the platform prior to the scheduled expiration of the job. Therefore, this number is an overestimate of the length of time that a job may be visible on the platform.

<sup>6</sup> Which users are exposed to the Jobs product in the News Feed is determined by an algorithmic system and is not dependent on the treatment assignment of the experiments discussed in this paper.

Figure 1: *Status quo* job search interface for Jobs on Facebook job-seekers



*Notes:* Interface shown to job-seekers on a mobile device.

by various treatments. As we can see in the figure, for each vacancy, the job-seeker could see the title of the job, whether it was part-time or full-time, the name of the employer, number of days since it was posted, and if the employer has posted the wage, the hourly wage. To learn more about the vacancy, the job-seeker had to click on the “tile” for that opening. Clicking exposed a “detailed view” of the job that included the full job description written by the employer. It also included an “apply” button that the job-seeker could use to submit an application.

### 2.1. Measurement of job-search behavior and vacancies

We now describe our measurement of vacancies and job-search behavior. We observe the job vacancies a user loaded onto the job board during their search, which is a function of how far they

scrolled down the device, their location, and their search query parameters. For each job vacancy, we observe the date it was posted, the date it was closed, and various meta-data both posted by the employer and inferred by algorithms. For example, employers specify job location, job title in the native language, and the opening type (full-time, part-time, contract, internship). Facebook also used algorithms to infer the Standard Occupational Classification (SOC) codes from the job posting. Over 70% of positions were full-time and the most common SOC codes were related to sales, driving, and fast food counter workers.<sup>7</sup>

We now consider the measurement of searcher behavior. A “view” occurred when a vacancy tile appeared on a screen. We also observe whether the user clicked on a vacancy to learn more, which we call a “detailed view.” Finally, we observe whether the job-seeker applied to a particular opening and some information about that application.

When job-seekers started an application, their information was populated into an application, using data from their Facebook profile—educational history, past employment, contact information and so on. Searchers could fill in additional information that is not already listed on Facebook. Our application measure is likely a lower bound on the number of job applications created, as in some cases, job-seekers would have enough information about the employer to apply directly.<sup>8</sup> However, the convenience of simply submitting through the Facebook App makes this the most likely course of action.

After an application occurs, we have imperfect information about what happened. For vacancies created through the JOF platform, which we call ‘native’, we can observe a variety of interactions including whether an employer viewed an application, whether the employer contacted an applicant through Facebook, and whether an employer told Facebook whether an interview was scheduled. Each step in this process is ‘leaky’, so that we see a large share of applications are viewed, but a much smaller share have contacts and interviews. This partially occurs because at each step employers and applicants can choose to take the interaction off of the platform. There are also vacancies which are syndicated from other platforms, which we refer to as ‘third-party vacancies’. In the US, third party platforms can be applicant tracking systems. For some of these vacancies, we cannot measure interactions between applicant and employer because the interactions take place off of the platform.

<sup>7</sup> Specifically, the top five 5-digit SOC codes are First-Line Supervisors of Sales Workers (41-101), Marketing and Sales Managers (11-202), Driver/Sales Workers and Truck Drivers (53-303), Sales Representatives, Wholesale and Manufacturing (41-401), Fast Food and Counter Workers (35-302). These comprise approximately 10% of all vacancies.

<sup>8</sup> There are some vacancies which do not allow the user to apply using the Facebook platform. These vacancies are syndicated from outside of JoF. We do not include these in our calculations of application rates.



### 3. Theory and evidence on the role of application order and vacancy age.

Job platforms are interested in helping their users find good matches. As part of this goal, they provide information to users in order to help them form these matches. In this section we describe a simple decision framework for what job seekers *should* care about. We then document two empirical regularities which motivate our experimental analysis of information about prior applications and vacancy age. First, that earlier applicants have a higher likelihood to have their applications viewed and responded to. Second, even conditional on the number of applications, applications sent to more recently posted vacancies are more likely to get a response. The advantages of applying earlier point to a role for information interventions.

#### 3.1. Decision framework

Job seekers would like to obtain the job that gives them the highest utility, which may stem from wages and other amenities. However, some jobs that would give a high utility cannot be obtained. This could be because the employer judges the worker unqualified for the job, because the competition for the job is too high, or because the employer has already interviewed the candidates who will be hired. Given these concerns, job seekers value information about the likelihood of obtaining a job in addition to the quality of the job.

Suppose job seekers (s) enter the market at a time (t), evaluate jobs (j) sequentially, and apply to a job if their expected benefit from applying exceeds the costs. Let a job seeker's expected utility from a job offer from vacancy j be  $u_{sjt}(z_s, x_j, \sigma_{tj})$ . This expected utility is a function of searcher characteristics,  $z_s$ , job characteristics,  $x_j$ , and information  $\sigma_{tj}$  about the vacancy's status at time t. Similarly, let a job seeker's belief about receiving and accepting an offer be  $b_{sjt}(z_s, x_j, \sigma_{tj})$ . The job seeker applies to vacancy j if:

$$b_{sjt}(z_s, x_j, \sigma_{tj}) * u_{sjt}(z_s, x_j, \sigma_{tj}) > c(n_s) \quad (1)$$

Note that the cost of applying  $c(n_s)$  may vary with the amount of prior applications sent by the searcher ( $n_s$ ). [Ursu et al. \(2022\)](#) show that consumers get tired or distracted from searching and that implies that search costs increase in the amount of time spent searching. These increasing search costs may explain why we observe a limited number of applications even in markets with many options and low costs of applying.

We are interested in understanding the effects of signals about a vacancy, such as the amount of competition and the age of the vacancy, affect the decision to apply. As we show in the next several sections, vacancies that have already received many applications are less likely to interview and hire new applicants. Similarly, vacancies that have been posted a longer time ago are less likely

to interview and hire. These stylized facts suggest that providing meta-information will lead job seekers to update their beliefs about an offer. As a result they should direct search effort to younger vacancies with few job applications.

In contrast, it could be that the number of prior applications is indicative of a higher utility job. For example, the fact that many others have applied could mean that those others have read the detailed job description and have found it appealing. Or that they know something about the employer that makes the application a particularly high value. In that case, information that a vacancy has many applications could actually increase the likelihood that job seekers apply.

We abstract away from the possibility that the information provided about a particular job conveys broader information about the platform as in [Tucker and Zhang \(2010\)](#) and [Fong \(2019\)](#). In [Section 5.4](#) we offer suggestive evidence to support this assumption. In particular, applications increase just for vacancies for which information is shown rather than for vacancies for which information is not shown.

### 3.2. Effects of applying earlier

We now show that applying earlier (whether measured by the application order or time) results in a higher probability that an application receives a positive outcome. The simplest explanation for this is that an employer checks applications at some time that the applicants cannot condition on. As a result, when everything else is held constant, earlier applications are more likely to be considered ([Van Ours and Ridder \(1992\)](#)).

To measure the effect of application order, we can compare the outcomes of multiple applications to the same vacancy. In particular, we estimate regressions of the following form:

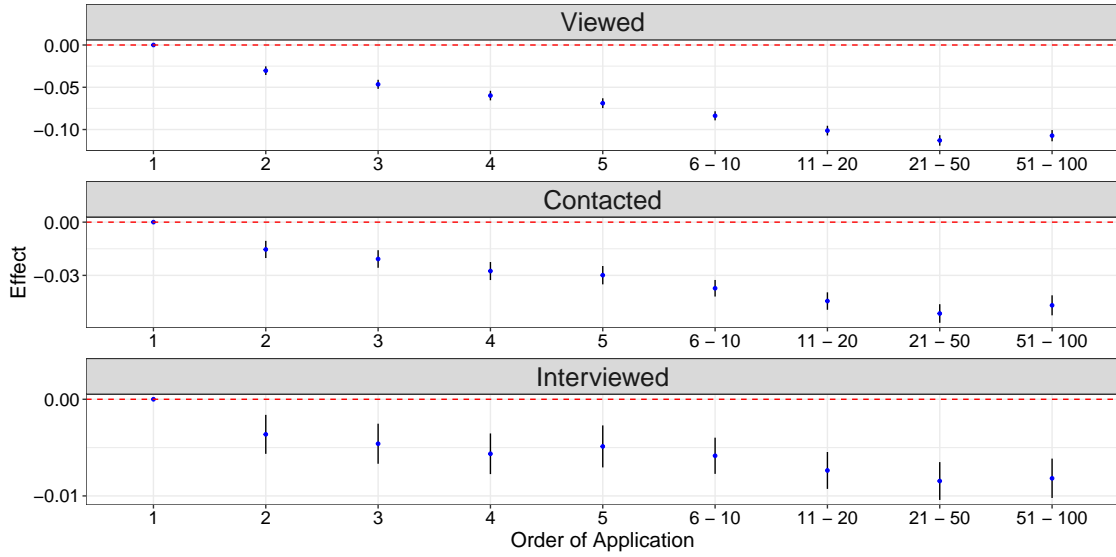
$$y_{i(s)j} = \beta_{\text{order},i(s)j} + \delta_{\text{age},i(s)j} + \kappa_j + \epsilon_{i(s)j} \quad (2)$$

where  $y_{i(s)j}$  are application outcomes for vacancy  $j$  by application  $i$  from searcher  $s$ ,  $\beta_{\text{order},i(s)j}$  are fixed effects for the order of the application (e.g. 1st application, 10th application),  $\delta_{\text{age},i(s)j}$  are fixed effects for the age of the vacancy at the time (e.g. 1 day, 3 days) at which the application was submitted, and  $\kappa_j$  are vacancy fixed effects. We estimate the above equation on a 5% sample of vacancies posted between March 3 and August 18 of 2019, the period during which our experiments ran. To ensure that the sample is balanced, we consider only the first 100 applications for vacancies that received at least 100 applications.<sup>9</sup>

[Figure 2](#) displays the estimated coefficients on application order from the above equation, where the first application is normalized to 0. Later applications have a lower probability of being viewed, contacted, and interviewed. To get a sense of the magnitudes involved, the baseline rate of interviews

<sup>9</sup> This keeps 27% of observations, see [Figure A.1](#) and [Figure A.2](#) for results with a full sample.

Figure 2: Relationship between application order on employer responses



*Notes:* Each point represents the estimated effect and each line present the 95% confidence interval for estimates of the effects of application order on whether an application is viewed, responded to, or results in an interview. The coefficient on the first application is normalized to 0. Standard errors are clustered at a vacancy level.

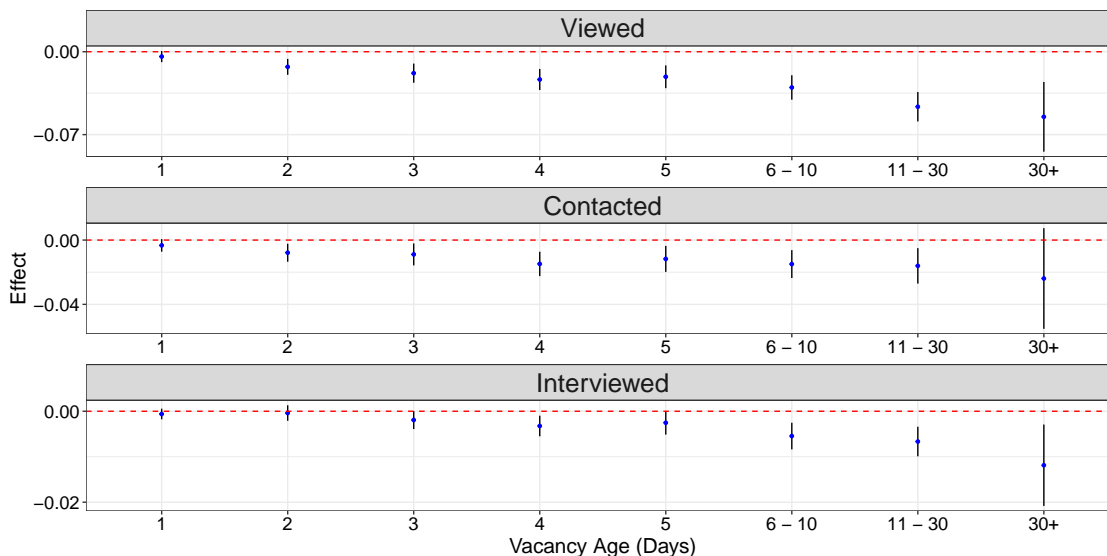
in this sample is 1.3%, which reflects the fact that many employers do not use the JOF interface to record interviews. We see that later applications experience a reduction in interview rates of almost 1 percentage point. To summarize, application order matters for application outcomes, and this justifies our focus on providing application order information to job-seekers.

We also see that most of the decline in application outcomes occurs by the fifth applications. This suggests that job seekers should care especially about being one of the first applicants. A theoretical justification for this can be seen as follows. Suppose that the vacancy has already received  $n$  applications and the job seeker is considering sending application  $n + 1$ . Furthermore, suppose that the vacancy is equally likely to hire each one and only hires one applicant. Then the hire rate is highest when there are few other applications and the derivative of the hire rate is highest at the second application and diminishes with the number of other applicants.

We also consider the effect of vacancy age on the likelihood of application success. We plot the estimated coefficients for vacancy age from Equation 2 in Figure 3, which include vacancy fixed effects and application order controls. We find that, even conditional on the order of the application, vacancy age is negatively correlated with employer views, contacts, and interviews. Although the coefficients are not as large in magnitude as those we found for application order, they are nonetheless statistically significant.

A related question is whether and how job-seekers know about the negative relationship between application order, vacancy age, and interview rates. This relationship is likely to hold for most vacancies, regardless of platform, so that anyone who's searched for a job before may have had a

Figure 3: Relationship between vacancy age at the time of application and employer responses



Notes: Each point represents the estimated effect and each line represents the 95% confidence interval for estimates of the effects of vacancy age at the time of application on whether an application is viewed, responded to, or results in an interview. The coefficient on the first application is normalized to 0. Standard errors clustered at a vacancy level.

chance to learn about it. Furthermore, many people have been on the hiring side of the market and could have observed that earlier applications get more attention. Lastly, people looking for advice online will find advice suggesting that earlier applications are more likely to be successful.<sup>10</sup>

#### 4. Experimental provision of information

Our primary research interest is in how information about the competition for a particular vacancy affects job-seeker decision-making. We study this question by analyzing three experiments conducted over a five month span in 2019.<sup>11</sup> The authors of this paper provided input into the design of these experiments but these experiments were primarily conducted for the purposes of improving the JOF product. The final decisions regarding which treatment arms to run and when were determined by product managers and designers. When large samples are possible, it is common practice for tech companies to test many minor variants of a design (see Kohavi et al. (2020) for details about large-scale experimentation at technology companies).

The experiments all varied the information job-seekers had about a vacancy when they viewed it. In total, there were 17 treatment arms across the experiments, which are summarized in Figure B.1. This is a lot of treatment arms relative to many academic experiments, but on platforms with vast

<sup>10</sup> For example, see this Quora question: “Does it make a difference if you apply for a job as soon as it is posted?”: <https://www.quora.com/Does-it-make-a-difference-if-you-apply-for-a-job-as-soon-as-it-is-posted>.

<sup>11</sup> Experiment I was conducted from 2019-03-26 to 2019-05-09 (44 days). Experiment II was conducted from 2019-05-31 to 2019-06-28 (28 days). Experiment III was conducted from 2019-07-22 to 2019-08-18 (27 days)

numbers of users, it is typical to try many minor variations of a treatment to determine the best one. A randomly chosen subset of  $\leq 50\%$  of job seekers were eligible for each of the experiments.<sup>12</sup> Across the experiments, treatment cells differed in (a) what information was displayed, (b) how spaced out the information was (with information only presented for a subset of tiles), (c) the color (grey vs blue) of the information (to make information more or less salient). Note that other aspects of Facebook’s systems, such as the ranking algorithm, did not use information about a user’s treatment assignment.

Figure 4 displays how the interface presented to job-seekers was altered by the treatments. Figure 4a displays one job tile when the number of applicants was less than 5. The color and the information varied depending on the treatment and the number of applications to the vacancy. Vacancies that had more than 4 prior applications had messages that display one of the following texts: ‘5 - 20 applications’, ‘21 - 50 applications’, ‘50 - 200 applications’, ‘201+ applications’. Figure 4b displays the control tile. Note that the control tile occupied *less* vertical space than the treatment tile. This will be important for our subsequent results given the limited screen space available on mobile devices. Figure 4c displays how each tile was combined in the JOF product.

Given the number of treatments available, we primarily analyze the experiment by pooling similar treatment arms. This allows us to simplify the exposition and increase our statistical power. Section B demonstrates that the specific manner in which a particular intervention was implemented within an experiment was not of first order importance to the treatment effect on applications. This appendix also demonstrates that the experiments are balanced on pre-treatment covariates—indicative of a successful randomization.

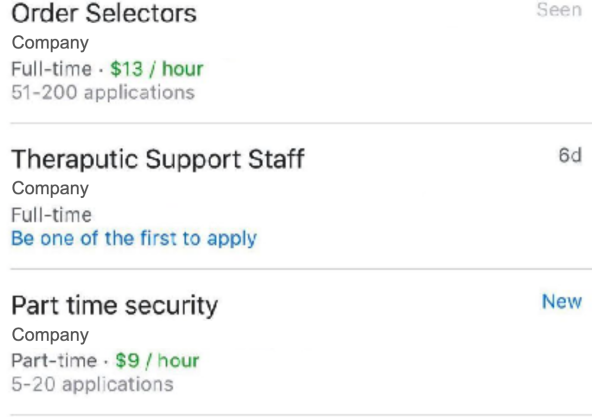
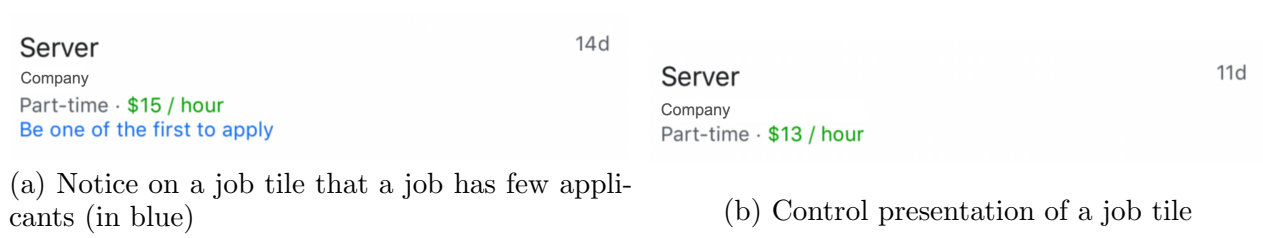
One concern with our experimental design is that there may be violations of the Stable Unit Treatment Value Assumption (SUTVA). In particular, when a treated searcher applies to a vacancy due to the treatment, this may affect the competition faced by the control searchers and may affect the job posting behavior of employers who receive applications induced by the treatment. As a result, our experimental analysis focuses on the effects on differences in individual behavior under the market conditions faced by searchers during our experiments. We are not able to capture how the treatment affects the equilibrium level of competition and match rates on the platform.

## 5. Effects of information about the number of prior applicants

Directly displaying information about the number of prior applicant increases applications and redistributes them towards relatively under-subscribed vacancies. The effect of the treatment on applications comes from information displayed about a particular vacancy.

<sup>12</sup> Note that for reasons of confidentiality the company did not permit us to report the exact percentage. Tech companies often allocate only part of the universe of users for an experiment in order to isolate the effects of potentially interacting concurrent experiments and in order to mitigate risk (Bakshy et al. (2014)).

Figure 4: Illustration of popularity information shown vacancy tiles



(c) Bin applicant count information

Notes: Job vacancy tile interfaces.

### 5.1. Overall job application intensity

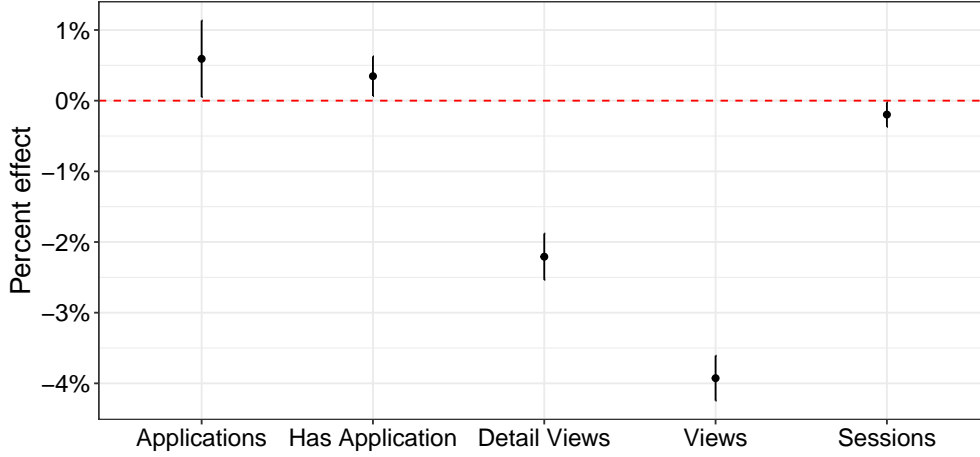
We begin by analyzing the aggregate job search effects of the pooled treatment before discussing its heterogeneous effects across vacancies. We estimate these effects by running regressions of the following form:

$$Y_s = \gamma_{exp} + \beta_1 Treat_s + \epsilon_s \quad (3)$$

where  $Y_s$  refers to outcomes for searcher,  $s$ , observed in the experiment and  $Treat_s$  refers to an indicator variable for whether the searcher was in the treatment group that provided information on the number of prior applicants. We also include fixed effects,  $\gamma_{exp}$ , for the experiment number (1 - 3). To get an effect in terms of percents, we take a ratio of  $\beta_1$  and the mean of  $Y$  in the control group. Figure 5 plots the treatment effects and standard errors<sup>13</sup> for the main variables of interest calculated across from the pooled sample consisting of 29,375,533 observations and Table B.1 displays the regression results in table form.

<sup>13</sup> Standard errors for this object are calculated via the delta method. We considered using randomization inference but the computational costs were high with our large sample size and the bias of the asymptotic standard errors is likely to be low with a large sample.

Figure 5: Effects of revealing competition information on job search behavior and outcomes



*Notes:* This figure plots the average effects in percent terms and their 95% confidence intervals. Each observation is a searcher in an experiment and all treatments that included information about prior applications were pooled. Standard errors are computed via the delta method. [Table B.1](#) displays the regression estimates used to generate this figure and [Figure A.3](#) displays the effects in levels.

Total applications increase by 0.59% and the share of searchers with at least one applications increases by 0.35%. This demonstrates that the treatment effects are coming from both the extensive and the intensive margin. In contrast to the effect on total applications, we find relatively large decreases in the number of views and detail views. These decreases in views are mostly a mechanical consequence of the fact that the information provided by the treatment takes up more space in the interface.<sup>14</sup> In [section B](#), we demonstrate this fact by showing that the treatment effect size is correlated with how frequently information is shown in a given treatment. The effect on the number of search sessions is less pronounced at -0.2%, which may explain why we nonetheless see increases in overall applications.

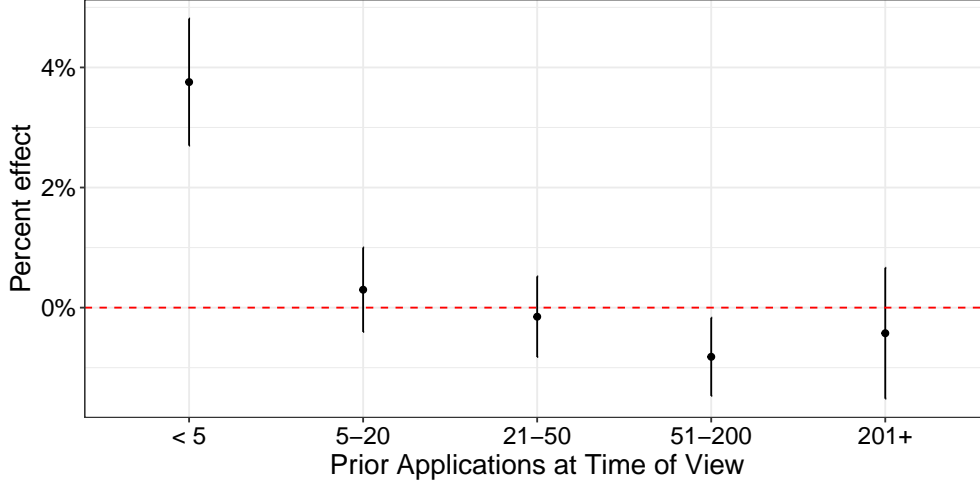
## 5.2. Evidence on competition avoidance

We now test whether job seekers respond to information by avoiding competition or herding. To do so, we consider the effect of the treatment on applications to jobs with differing amounts of prior applications. We find that the largest increases in applications occur for vacancies with few prior applications and that there are negative treatment effects for vacancies with many prior applications. Our results favor the competition avoidance mechanism and not the herding mechanism.

Our estimation strategy follows [Equation 3](#) with the outcome,  $Y_{s,b}$ , equal to the number of applications sent by seeker  $s$  to jobs with a number of applications at time of exposure in a

<sup>14</sup> To see that the negative effect on views is plausible, suppose that a mobile phone screen can fit four vacancies on average, and that each vacancy takes up three lines without the treatment. Adding one extra line takes up an extra  $1 - (4 \times 3) / (4 \times 3 + 1) = 7.7\%$  space. Given the diversity of mobile phones and differences in search activity, it is plausible that this extra line can affect whether a vacancy is viewed.

Figure 6: Effects of competition information on applications to different status vacancies



*Notes:* This figure plots the average effects in percent terms and their 95% confidence intervals. Each observation is a searcher in an experiment and all treatments that included information about prior applications were pooled. For each searcher, we observe the number of applications in the experiment sent to jobs with a number of prior applications in a given bin. This calculation is done at the time the seeker is first exposed to a particular vacancy. Standard errors are computed via the delta method. [Table B.2](#) displays the regression estimates used to generate this figure.

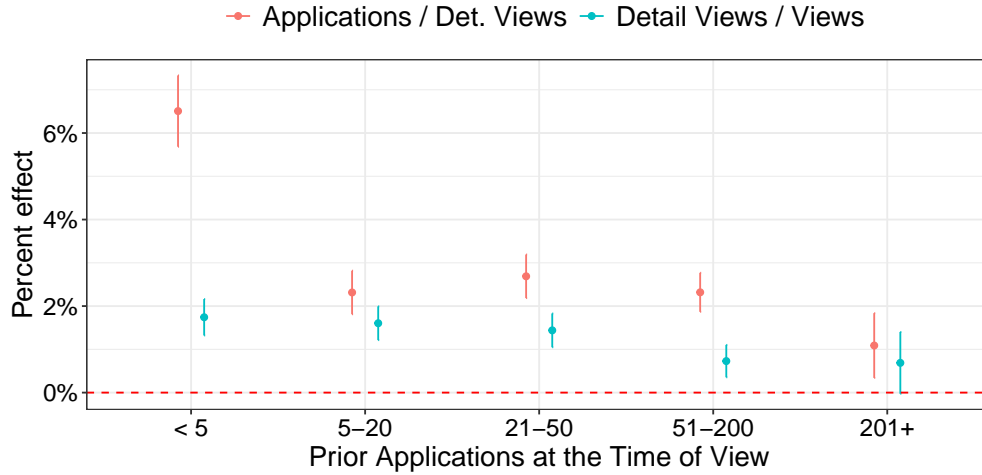
particular range. For example, if a job seeker sends two applications to vacancies that had fewer than five applications at the time,  $Y_{s,<5} = 2$ . We bin outcomes in a way that parallels the information treatment. [Figure 6](#) displays the results.

A striking pattern in [Figure 6](#) is that the effect on applications is largest for vacancies with fewer than five applications. This is consistent with our results that there is a particularly large benefit to being one of the first few applications ([Figure 2](#)) and with results in [Van Ours and Ridder \(1992\)](#). The effect on competition information is not driven by the frequency with which information is shown or by whether information about other application bins is shown. [Figure B.5](#) shows that the effect on the  $< 5$  category is similarly sized for seven different treatment arms. In [Figure C.3](#), we investigate heterogeneity by a number of factors including gender, age, and device and fail to detect statistically significant differences.

We can also bring our empirical specification closer to that of [Gee \(2019\)](#) to make a direct comparison. [Gee \(2019\)](#) considers applications per detail view and finds that information about the number of applicants increased applications per view by 3.5% but that that effect did not vary in a systematic manner by the exact information shown. In [Figure 7](#), we display the effects of the treatment on the applications per view and views per application. Conditional on viewing a vacancy, treatment users are more likely to apply to a vacancy. [Figure 7](#) shows that this effect



Figure 7: Effects on detail views / views by prior applications to a vacancy



Notes: This figure plots the average effects in percent terms and their 95% confidence intervals. The outcome is the ratio of detail views to views. Each observation is a searcher in an experiment with at least one view or detail view for a job in the bin and all treatments that included information about prior applications were pooled. Standard errors are computed via the delta method.

is especially large for vacancies with fewer than five prior applications. This supports our theory that searchers avoid competition.<sup>15</sup>

### 5.3. Survey evidence on competition avoidance

Our field experimental results show that, on average, job seekers prefer to apply to jobs with few other applicants. However, these results tell us little about heterogeneity in these preferences or causal mechanisms. To further investigate heterogeneity and causal mechanisms, we designed a survey choice experiment. This experiment corroborates our prior results that the vast majority of job seekers avoid competition due to application about the number of applicants. Herding is not quantitatively important.

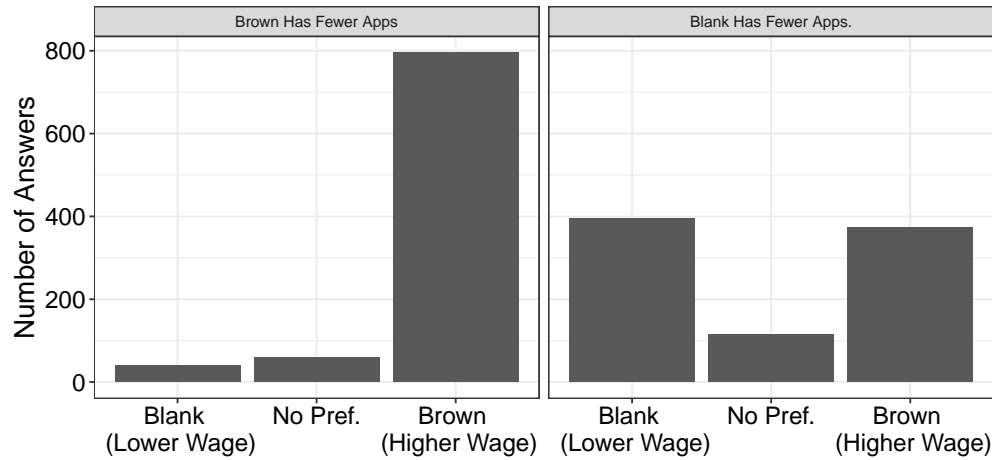
The pre-registered survey experiment<sup>16</sup> is designed as follows. Participants recruited from the online platform Prolific are faced with three choice scenarios (see Figure D.1). In each scenario, they must choose the job they value applying to the most if they were unemployed. The jobs differ in their wages and current applications. There were a total of 592 participants in the sample comprising the main analysis, and each participant answered three choice questions.

Figure 8 shows that participants were much more likely to choose the lower wage option when it has fewer current applicants. Almost no-one choose the lower wage option when it had more

<sup>15</sup> The overall increase in the treatment ratios of detail views to views can be explained by the fact that the treatment reduced views and detail views.

<sup>16</sup> The experiment was determined to be exempt from the IRB by MIT's Committee on the Use of Humans as Experimental Subjects. The preregistration for the experiment is available here: <https://www.socialscienceregistry.org/trials/9344>.

Figure 8: Distribution of survey responses



*Notes:* This figure plots the distribution of responses in the choice survey. The left side contains responses for which Brown Co is displayed as having fewer applications and having a higher wage. The right side contains responses for which Blank Co has fewer responses and a lower wage.

current applicants. Section D contains additional analysis of the survey, including analysis of textual responses, regression analysis, and heterogeneity analysis by gender and recent job search experience.

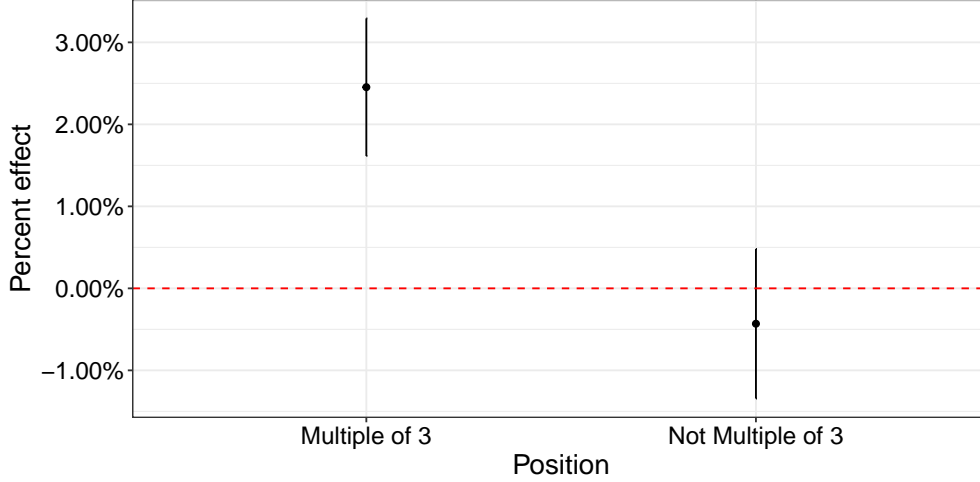
#### 5.4. Spillovers of competition information

We now consider whether the effect of the treatment comes solely from the information about a particular vacancy or whether there are spillovers onto vacancies for which no information is provided. A positive signal about a particular vacancy may draw away applications from other vacancies or it may induce additional applications to other positions due to learning. We can study these mechanisms by considering treatments which display information every three tiles rather than every tile. If information causes substitution, then we should see that the treatment has negative effects on vacancies that appear on tiles that are not multiples of three. On the other hand, if there is learning, we should see positive effects for these tiles.

Figure 9 plots the treatment effects on applications based on the position in which they were shown. The estimates are pooled across two treatment arms for which information is displayed every third tile and only when the vacancy has fewer than five applications. We see a positive and statistically significant effect for vacancies in a position divisible by 3. In contrast, for vacancies in other positions, we see a negative but small in magnitude and insignificant effect. This coefficient is consistent with some level of negative spillovers between ads with and without information, although if we take the point estimates at face value, then the negative spillovers on the two positions without information ( $2 \times 0.5\%$ ) are smaller than the positive effect on the vacancy with the information ( $2.5\%$ ). As a result, it's likely that the applications induced by the treatment

come from the information learned about particular vacancies and that negative spillovers to other vacancies do not wholly outweigh the benefits to the treated vacancy.

Figure 9: Effects of competition information on applications, by position of vacancy



*Notes:* This figure plots the average effects in percent terms and their 95% confidence intervals. The outcome is the number of applications. Each observation is a searcher in a treatment arm where information is displayed every 3 tiles or in the control. Standard errors are computed via the delta method.

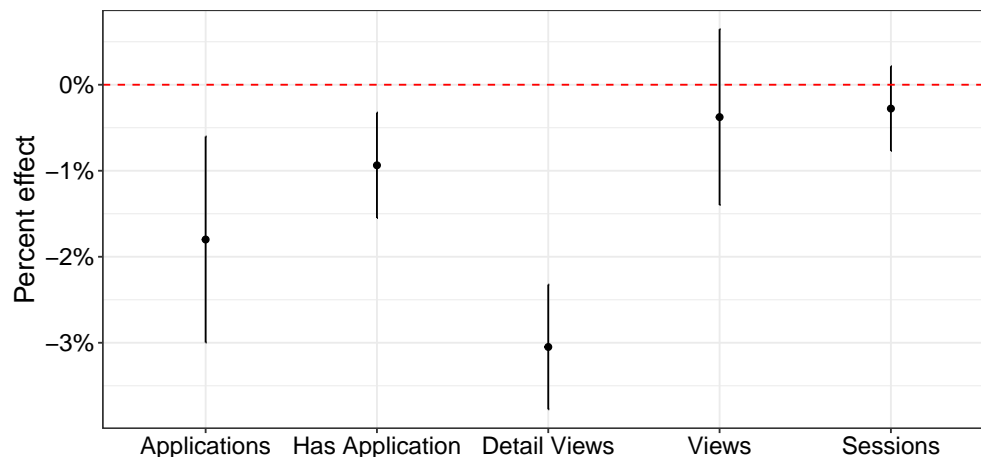
## 6. Vacancy age effects

To summarize, we have shown that job seekers value information about the level of competition when applying for vacancies and respond by applying to vacancies with fewer prior applications. We now investigate whether searchers use other information as a proxy for vacancy competition. Job-seekers often know the age of a vacancy because it is provided directly by the platform and JOF provided this information by default. To understand how this information is used, we study how job-seekers without access to this vacancy age information responded. We find that job searchers use the vacancy age as a proxy for competition when other information is not available and that, when vacancy age information exists, competition information helps to attract applicants to older vacancies.

Our empirical strategy is to study the effects of a treatment arm in which vacancy age is not displayed but everything else is held constant. This was the case in one treatment arm of Experiment I. In the control group, the tile displayed ‘New’ in blue if the vacancy was 5 or fewer days old, and would display ‘xd’ in grey otherwise, where ‘x’ is the number of days the vacancy has been posted (See [Figure 1](#)).

[Figure 10](#) displays the overall effects of removing vacancy age. We see that treated users submitted fewer applications, were less likely to apply to any job, and clicked on fewer vacancies. These effects are of comparable magnitude to the effects of including information on prior applicants.

Figure 10: Treatment effects on job search outcome  
Removing vacancy age



*Notes:* This figure plots the average effects in percent terms and their 95% confidence intervals. Each observation is a searcher in experiment I who was either in the control group or in the treatment group for which vacancy age was removed.

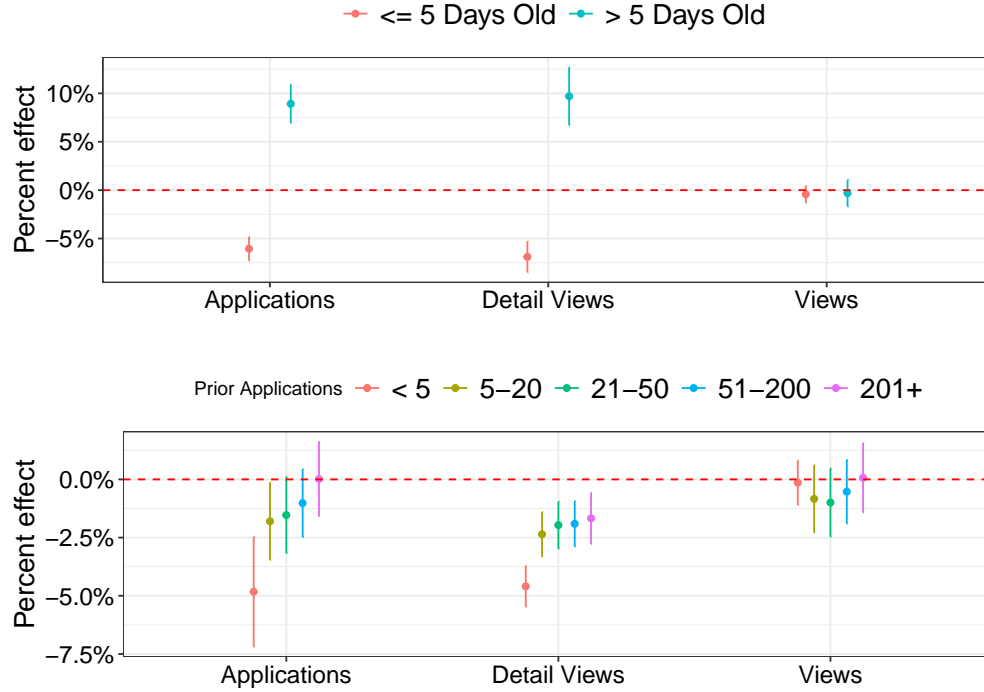
We further reinvestigate the role of vacancy age signals by calculating treatment effects based on the age of the vacancy at the time of the view. The top panel of [Figure 11](#) displays the treatment effects split by whether the vacancy had the text ‘New’ or not. We see that removing vacancy age decreased applications and detail views to new vacancies and increased them for vacancies older than five days. This heterogeneous effect suggests that users prefer applying to new vacancies when both new and old vacancies are identifiable, but otherwise cannot perfectly predict vacancy age based on observed information.<sup>17</sup>

Since vacancy age is correlated with the number of prior applications, we consider whether vacancy age information helps direct searchers to vacancies with less competition. The lower panel of [Figure 11](#) displays how removing vacancy age affects applications to vacancies with different levels of competition. The treatment causes the largest falls in applications for vacancies with fewer than 5 prior applicants and it has no effects on vacancies that receive over 200 applications.

If vacancy age already allowed searchers to find low competition vacancies, then why did the information treatments have an effect? One reason may be that this information allows seekers to identify older vacancies with little competition. [Figure A.4](#) shows the treatment effects of competition information for applications to vacancies with different ages. We find that the treatment increased applications to older vacancies, confirming this conjecture.

<sup>17</sup> Seekers may also try to infer vacancy age from other job characteristics or the ranking of the result. In this sense, our estimates represent a lower bound on the effects of vacancy age on search.

Figure 11: Treatment effects by job type  
Removing vacancy age



*Notes:* This figure plots the average effects in percent terms and their 95% confidence intervals. Each observation is a searcher in experiment I who was either in the control group or in the treatment group for which vacancy age was removed.

## 7. Effects on application outcomes

Signals of vacancy competition help searchers direct their applications to vacancies with less competition. However, whether this redirection of search is good for searchers or the platform is unclear. We now study this question by comparing application outcomes submitted across treated and control users.

Applicants who were shown competition signals experience a lower degree of realized competition for their applications. Column 1 of [Table 1](#) reports the results of a regression of the log of the application order on the pooled treatment indicator. We see that the order of treated applications was 1.5% lower than the order control applications. This difference in application order could be due to selection — namely changes in the types of vacancies applied to — or changes in the speed with which applications are sent. In column 2, we report results from the same regression but with vacancy fixed effects. Conditional on a vacancy, treated applications were submitted .2% faster. The conditional effect is much smaller than the unconditional effect — demonstrating that most of the reduction in competition is due to redirecting applications to vacancies with fewer other competitors.

Table 1: Application order and characteristics

	Log App Order (1)	Log App Order (2)	Log Eventual Apps (3)	Third Party Vac. (4)	Viewed (5)
Treatment	-0.0147*** (0.0013)	-0.0021** (0.0007)	-0.0127*** (0.0012)	0.0017*** (0.0002)	
Log App Order					-0.0777*** ( $6.78 \times 10^{-5}$ )
R <sup>2</sup>	0.040	0.834	0.047	0.006	0.125
Observations	13,846,468	13,846,468	13,846,468	13,846,468	12,765,466
Experiment fixed effects	✓	✓	✓	✓	✓
Vacancy fixed effects		✓			

*Notes:* This table contains results for a linear regression of applications outcomes on treatment (competition information), where all applications sent in the experimental sample are observations. ‘Order’ refers to the order in which the application arrived and ‘Eventual’ is the cumulative applications ever received by a vacancy. ‘Third party’ refers to a vacancy syndicated from a third party platform. ‘Viewed’ refers to whether the application was viewed by the employer.

We investigate this selection effect further in columns 3 and 4. Column 3 shows that treated applications are sent to vacancies that receive fewer eventual applications. Column 4 shows that these applications are also more likely to be sent to third party applications, showing that information about competition is especially important for these vacancies.

We now investigate whether the treatment induces better outcomes. There are several ways in which the treatment could have affected outcomes. First, since treated applications arrived earlier, they should have had higher success rates, all else equal. Second, since treated applications went to different vacancies on average, these vacancies could have had different proclivities to hire or not. Lastly, since application rates changed, treated applicants may have sent their applications to better or worse matching vacancies. The sign of these effects is theoretically ambiguous.

The treatment had negligible effects on the success rate of applications. [Table 2](#) displays outcomes across treated and control applications for applications sent to vacancies for which we can measure outcomes. Column 1 demonstrates that treated applications to vacancies were more likely to be viewed by the employers. The magnitude of this effect is .18%, which is very small. We also consider the effects of the treatment on interviews and hires — outcomes which more directly related to what the applicant cares about. Columns 2 and 3 of [Table 2](#) show tiny and not statistically significant effects on these outcomes. These effects are precisely close to zero — the 95% confidence interval excludes effects on the order of more than .25% ( $\frac{-.0003-1.96*.0003}{.273}$ ) in magnitude for contacts by employers.

To summarize, we find that treated applications face less competition, and that this is a function of both applying to different vacancies and to applying to applying earlier. We find precisely small effects on application outcomes. These two results combined suggest that the treatment was

Table 2: Differences in application outcomes

	Viewed (1)	Contact (2)	Interview (3)
Treatment	0.0008* (0.0004)	-0.0003 (0.0003)	$-2.09 \times 10^{-5}$ ( $9.46 \times 10^{-5}$ )
Mean of Y:	0.455	0.273	0.017
R <sup>2</sup>	0.052	0.046	0.001
Observations	12,765,466	12,765,466	12,765,466
Experiment fixed effects	✓	✓	✓

*Notes:* This table contains results for a linear regression of applications outcomes on the treatment in experiment 3. ‘Viewed’ is an indicator whether the employer viewed the application, ‘Contact’ is an indicator for whether an employer sent an applicant a message, and ‘Interview’ is an indicator for whether an employer marked that an interview was conducted.

successful in accomplishing the platform’s goals. Treated users did not experience worse outcomes, and at the same time, applied more often and especially to less competitive vacancies.

## 8. Conclusion

Social signals are used across platforms to influence search behavior and other activity. We’ve investigated the role of social and other signals in the context of job search, where the sign of the effect of social influence is ex-ante uncertain. On the one hand, telling job seekers that few people have applied could be a signal that the job is low quality. On the other hand, this information can also convey that the job seeker has a higher chance of getting the job. We find that the later effect dominates. Job seekers prefer to apply to jobs with very few prior applicants.

Even in the absence of competition information, applicants were able to direct their search towards less competitive vacancies. In a complementary experiment, we find that the job seekers strongly preferred new vacancies when information on vacancy age was available. Information about vacancy age greatly increased applications on the platform and redirected those applications away from popular but old vacancies. Both vacancy age and competition information increase usage of the platform, and did not harm application outcomes. These results point to the positive effects of displaying this information.

We studied just a few of the many information design decisions by the platform. For example, the platform could change where and how information is displayed. The platform could also create better signals of competition and match quality and display these signals to job seekers. Information design decisions may also interact in important ways with other platform design decisions such as ranking and user acquisition strategies.

The welfare implications of information design decisions are difficult to study with user level experiments. A key reason for this difficulty is that actions by seekers and employers exert externalities on each other, meaning that treatment effects from a user level experiment may not be

indicative of market outcomes were everyone to be treated. Another difficulty is that the platform has very noisy information about hiring and the quality of those hires. This means there is no direct measure of welfare, either for the searcher or for the employer. Lastly, the level of competition per application is an equilibrium object and the treatment may affect this equilibrium in a manner which our experiment is poorly suited to studying. Future work may be able to address these limitations with commuting zone level experiments, better measurement, and structural models of job search in digital platforms.

Finally, no single platform has a bird's eye view of the entire labor market. Both searchers and employers multi-home across a variety of platforms. As a result, measures of competition on one platform may not fully reflect the true level of competition and optimizations made on one platform might not improve outcomes in the entire labor market. The implications of this fragmentation in labor market platforms are important to understand for market designers and policymakers.

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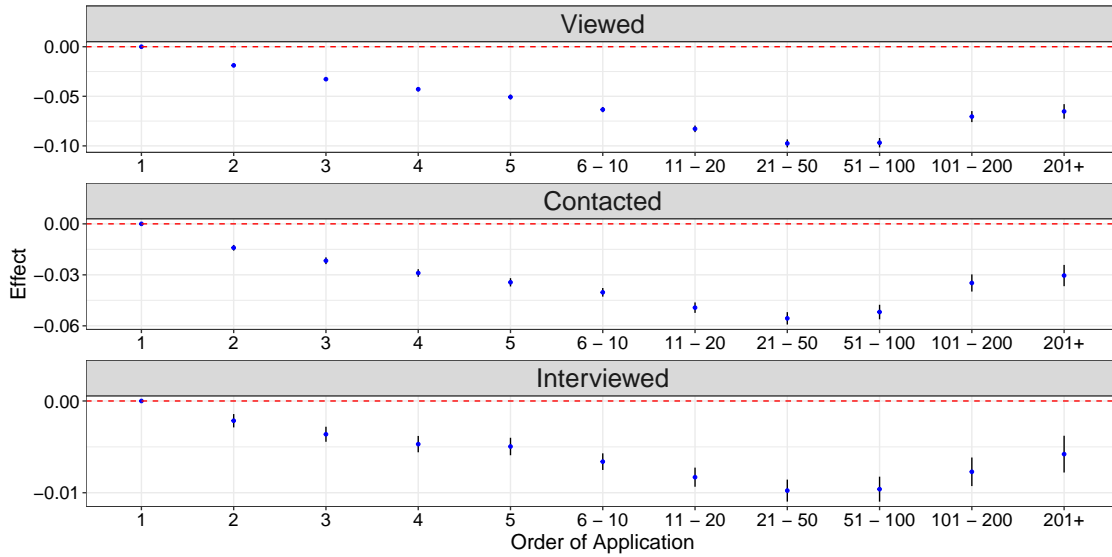


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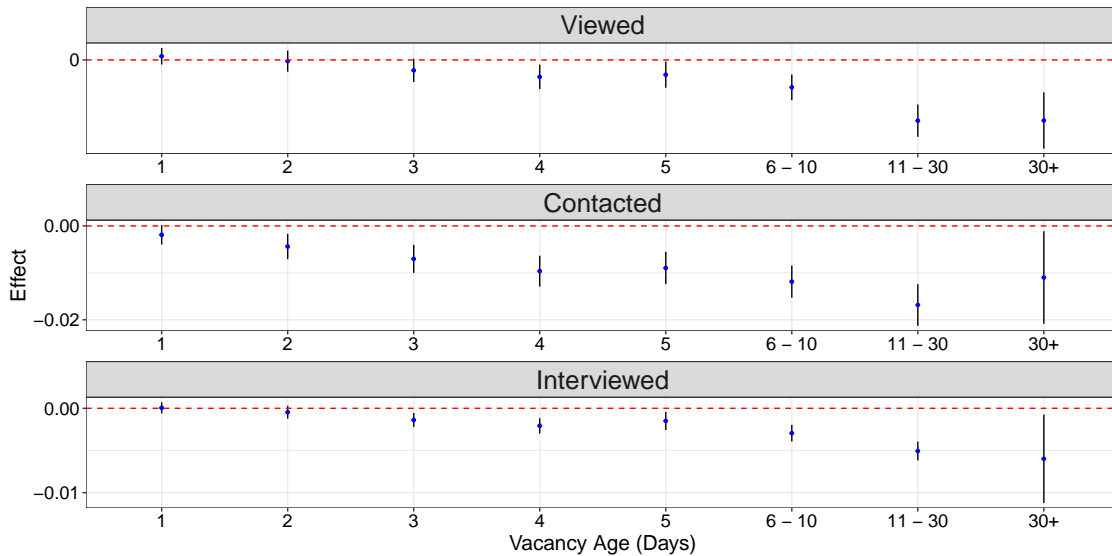
## Appendix A: Additional Figures and Tables

Figure A.1: Effects of application order - full sample



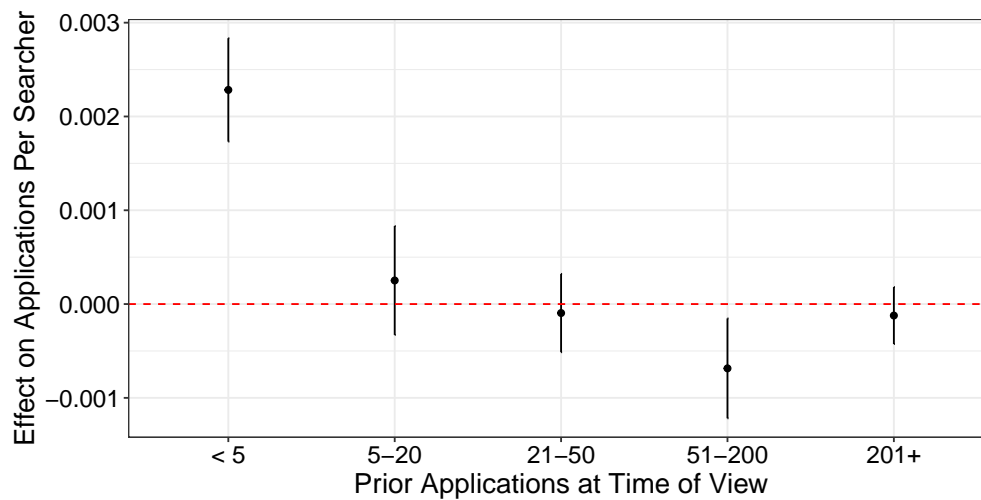
Notes: Each point represents the estimated effect and each line present the 95% confidence interval for estimates of the effects of application order on whether an application is viewed, responded to, or results in an interview. The coefficient on the first application is normalized to 0. Standard errors are clustered at a vacancy level.

Figure A.2: Effects of application order - full sample



Notes: Each point represents the estimated effect and each line represents the 95% confidence interval for estimates of the effects of vacancy age at the time of application on whether an application is viewed, responded to, or results in an interview. The coefficient on the first application is normalized to 0. Standard errors are clustered at a vacancy level.

Figure A.3: Effects (in levels) of competition information on applications to different status vacancies



*Notes:* This figure plots the average effect of the treatment (pooled across all arms) in levels. Each observation is a searcher in an experiment and all treatments that included information about prior applications were pooled.

Figure A.4: Treatment effects on vacancy age of applications  
Competition information



*Notes:* This figure plots the average effects in percent terms and their 95% confidence intervals. Each observation is a searcher in the experiments who was either in the control group or in the treatment group for which competition information was shown.

## Appendix B: Analysis by Experimental Arm

In this section, we describe and analyze the treatment arms of each of the three experiments. We begin with by describing the set of treatments used in the study. Figure B.1 shows how treatment parameters varied across arms and experiments.

The first dimension along which treatments differed was in whether every vacancy tile was eligible to show competition information. Column 1 contains the set of arms where information could be shown on every tile, while columns 2 and 3 contain arms where information could be shown either every 3 tiles or every 10 tiles (beginning with the first tile on the screen). Next, row 1 displays the set of treatments where information about competition was shown only for vacancies that had fewer than 5 prior applications. For these vacancies, the text ‘Be one of the first to apply’ was displayed.<sup>18</sup> Row 2 displays the set of treatment arms for which competition information could also be shown for vacancies with more than 4 applications. For these vacancies, the following text could be shown, where appropriate: ‘Be one of the first to apply’, ‘5 - 20 applications’, ‘21 - 50 applications’, ‘50 - 200 applications’, ‘201+ applications’. Treatment arms also differed by whether they displayed this information in blue (vs grey) always (‘All’), just for vacancies with  $< 5$  applications (‘First’), or never (‘None’). Finally, the grid excludes one treatment arm from Experiment 3, in which some signals were eligible to be shown every tile, while those relating to vacancies with  $< 5$  vacancies could only be shown on every third tile.

To check that the randomization was properly conducted, we preformed a set of balance tests. Figure B.2 displays these tests, where the p-value for the difference in means on pre-treatment covariates is displayed for every treatment arm in each experiment. Across four covariates (Age, Android User, Gender, and US user), we find differences that are not statistically significant at a 5% p-value. This evidence suggests a proper randomization of the treatment arms by Facebook in each experiment.

In addition to treatments with social information, Experiment 1 also contained arms that varied whether vacancy age was displayed. One of these arms was discussed in Section 6. Two other arms removed vacancy age, but added competition signals (either just ‘Be the first to apply’ or all competition signals).

<sup>18</sup> The text was also translated into the appropriate language for each locale.

Figure B.1: Treatment arms for experiments

	Every Tile	Every 3 Tiles	Every 10 Tiles																								
Only "Be one of the first to apply"	<table><tr><td>Exp.</td><td>Blue</td></tr><tr><td>1</td><td>None</td></tr></table>	Exp.	Blue	1	None	<table><tr><td>Exp.</td><td>Blue</td></tr><tr><td>2</td><td>All</td></tr><tr><td>3</td><td>All</td></tr></table>	Exp.	Blue	2	All	3	All	<table><tr><td>Exp.</td><td>Blue</td></tr><tr><td>2</td><td>All</td></tr></table>	Exp.	Blue	2	All										
Exp.	Blue																										
1	None																										
Exp.	Blue																										
2	All																										
3	All																										
Exp.	Blue																										
2	All																										
All Congestion Signals*	<table><tr><td>Exp.</td><td>Blue</td></tr><tr><td>1</td><td>None</td></tr><tr><td>3</td><td>First</td></tr></table>	Exp.	Blue	1	None	3	First	<table><tr><td>Exp.</td><td>Blue</td></tr><tr><td>2</td><td>All</td></tr><tr><td>2</td><td>None</td></tr><tr><td>2</td><td>First</td></tr><tr><td>3</td><td>First</td></tr></table>	Exp.	Blue	2	All	2	None	2	First	3	First	<table><tr><td>Exp.</td><td>Blue</td></tr><tr><td>2</td><td>All</td></tr><tr><td>2</td><td>None</td></tr><tr><td>2</td><td>First</td></tr></table>	Exp.	Blue	2	All	2	None	2	First
Exp.	Blue																										
1	None																										
3	First																										
Exp.	Blue																										
2	All																										
2	None																										
2	First																										
3	First																										
Exp.	Blue																										
2	All																										
2	None																										
2	First																										

\*Experiment 3 also had a treatment arm that displayed 'Be one of the first to apply' on only the 3rd tile when eligible but displayed other congestion information in grey on every tile when eligible.

*Notes:* This figure displays the experiments during which each combination of treatments appeared. Information was presented either every tile, every 3 tiles (starting with tile 1), or every 10 tiles (starting with tile 10). Treatment arms varied by whether only under-subscribed vacancies (< 5 prior applications) were marked with competition information, or whether all eligible vacancies were marked with competition information. Lastly, in certain cases competition information was given a blue color. Values of 'First' in the 'Blue' column denote that only signals for under-subscribed vacancies were given a blue color.

Figure B.2: Covariate balance test p-values across experiments

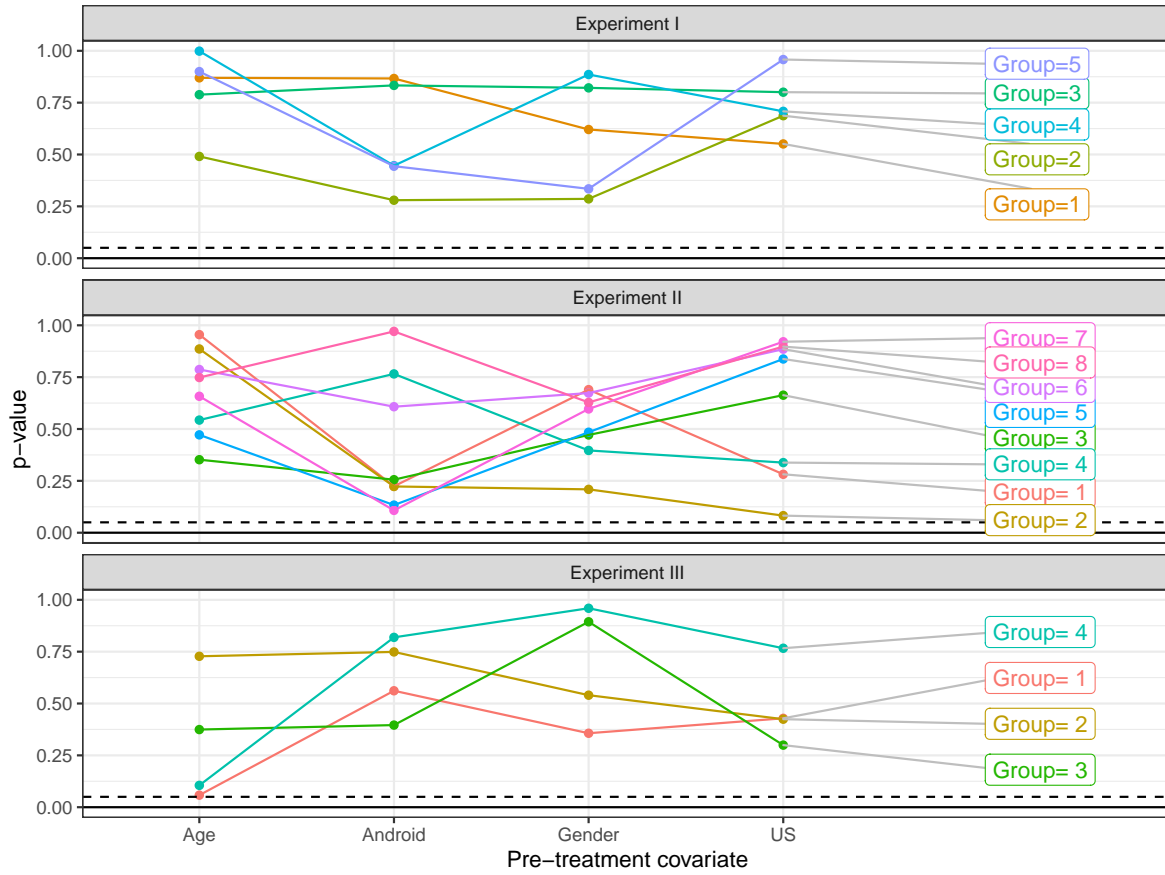
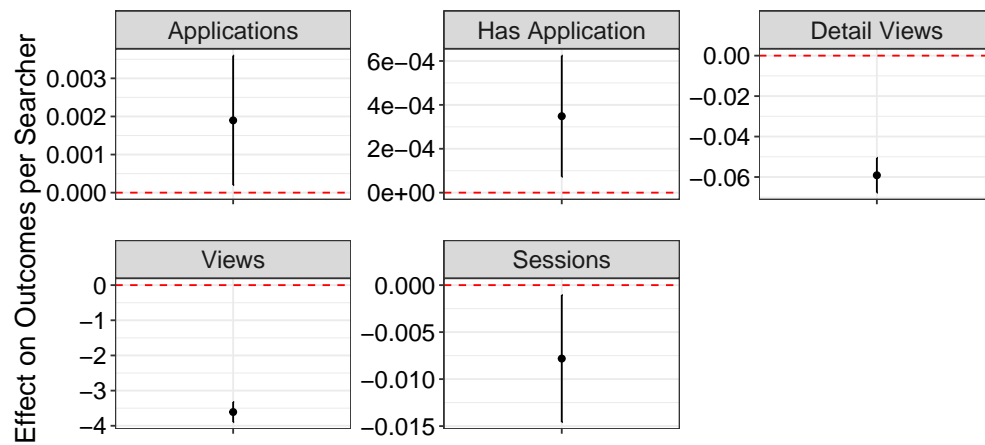
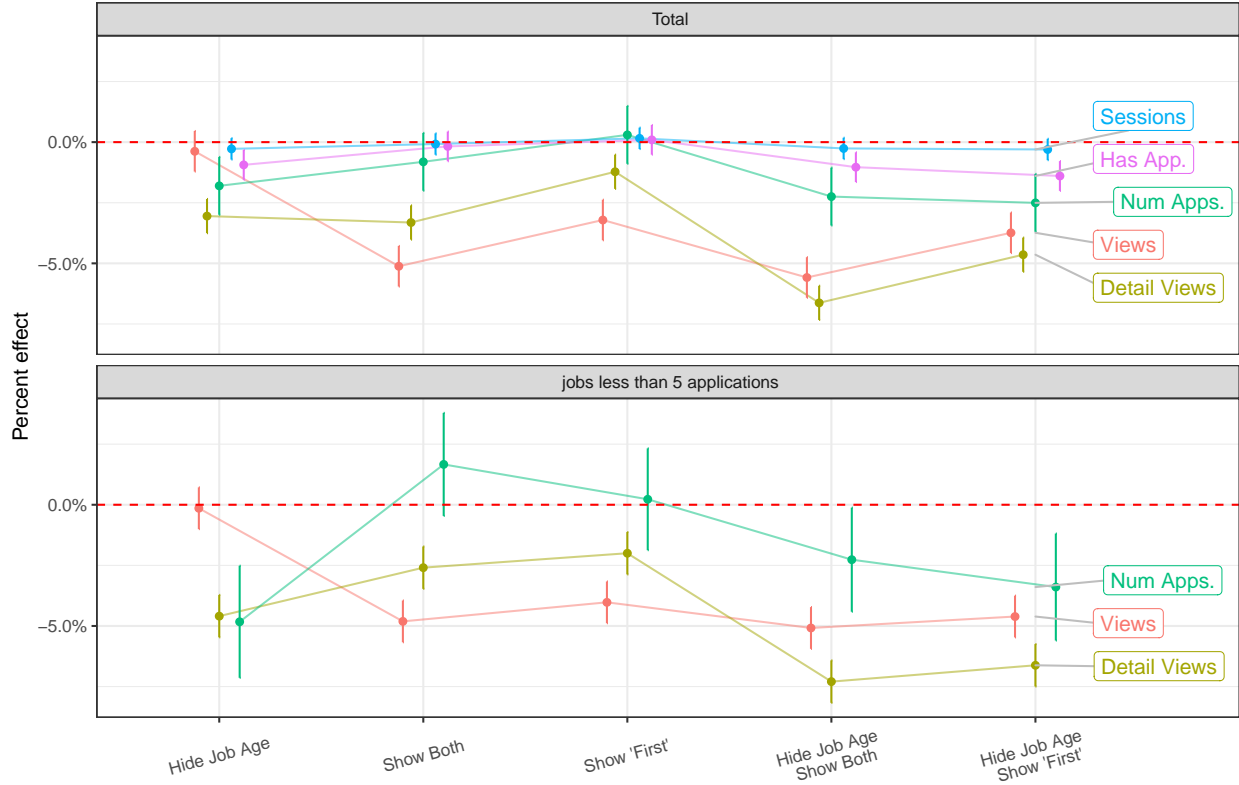


Figure B.3: Effects (in levels) of revealing competition information on job search behavior and outcomes



*Notes:* This figure plots the average effect of the treatment (pooled across all arms) in levels. Each observation is a searcher in an experiment and all treatments that included information about prior applications were pooled.

Figure B.4: Treatment Effects for Experiment 1



### B.1. Effects by Treatment Arm

Next, we discuss the by-arm treatment effects for each treatment and experiment. We begin with Experiment 1 (Figure B.4). Columns 1, 4, and 5 of the figure plot the treatment effects where the vacancy age is hidden. Columns 2 - 5 plot treatments where competition information is added. Broadly, the treatments where vacancy age is hidden experience drops in views, detail views and applications. Columns 2 and 3, where competition information is added but vacancy age remains. The two treatments have similar effects on our outcomes.

Figure B.5 displays the effects of the separate treatment arms of experiment 2. Broadly, the effects are of similar magnitude across arms. The clearest difference is that there is a bigger drop in views when competition information is displayed every 3 tiles rather than every 10 tiles. This drop is expected since the competition information takes up an additional line of text and therefore fewer vacancies can be shown in the 'every 3' treatments.

Finally, Figure B.6 displays the effects of the separate treatment arms of experiment 3. As in the other experiments, the effects on applications and sessions are similar across treatment arms. As before, the more frequently competition information is shown, the fewer vacancies are seen by the searchers.



Figure B.5: Treatment effects for experiment 2

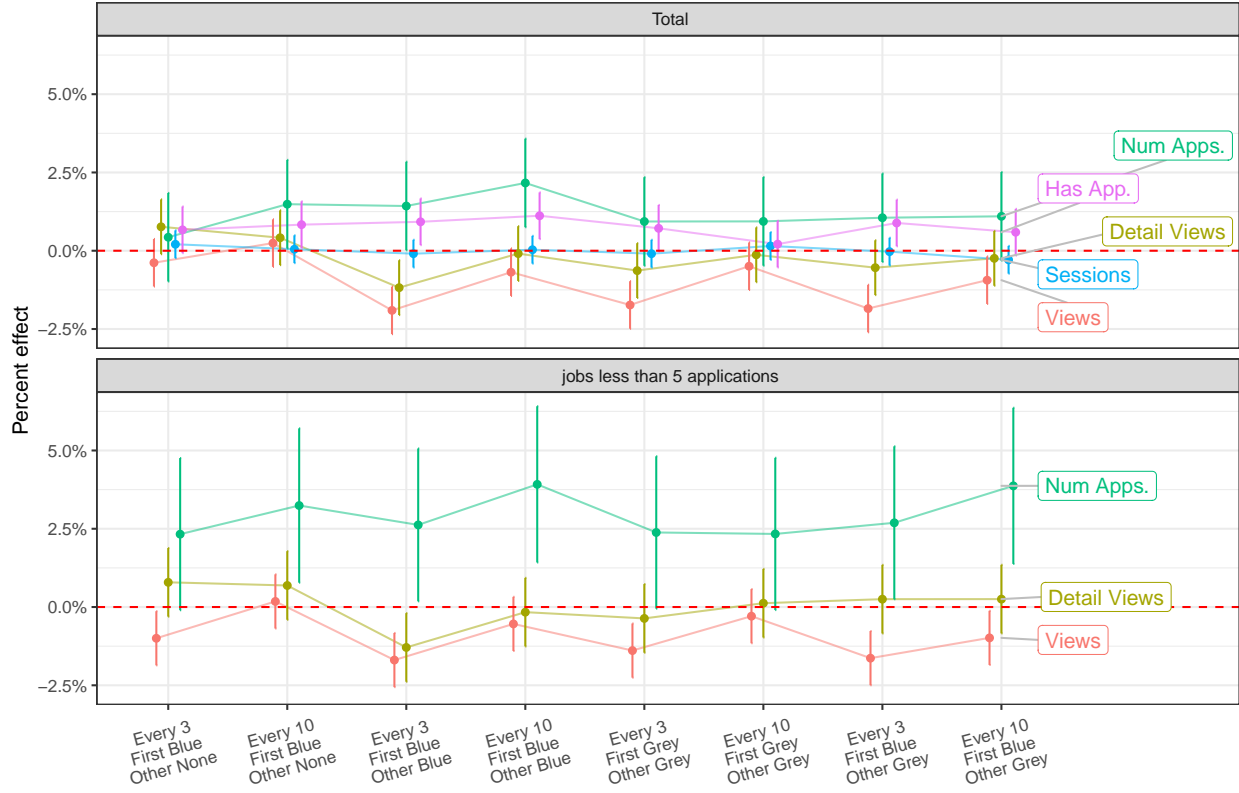


Table B.1: Treatment effects pooled across all three experiments

	Num. App. (1)	Has App. (2)	Detail Views (3)	Views (4)	Sessions (5)
Treatment	0.0019* (0.0009)	0.0003* (0.0001)	-0.0591*** (0.0044)	-3.609*** (0.1463)	-0.0078* (0.0034)
Mean of Y:	0.332	0.105	2.727	92.455	4.162
R <sup>2</sup>	0.002	0.007	0.002	0.001	0.008
Observations	29,375,533	29,375,533	29,375,533	29,375,533	29,375,533
Experiment fixed effects	✓	✓	✓	✓	✓

Notes: Effects of the competition signal treatment pooled across experiment.

Figure B.6: Treatment effects for experiment 3

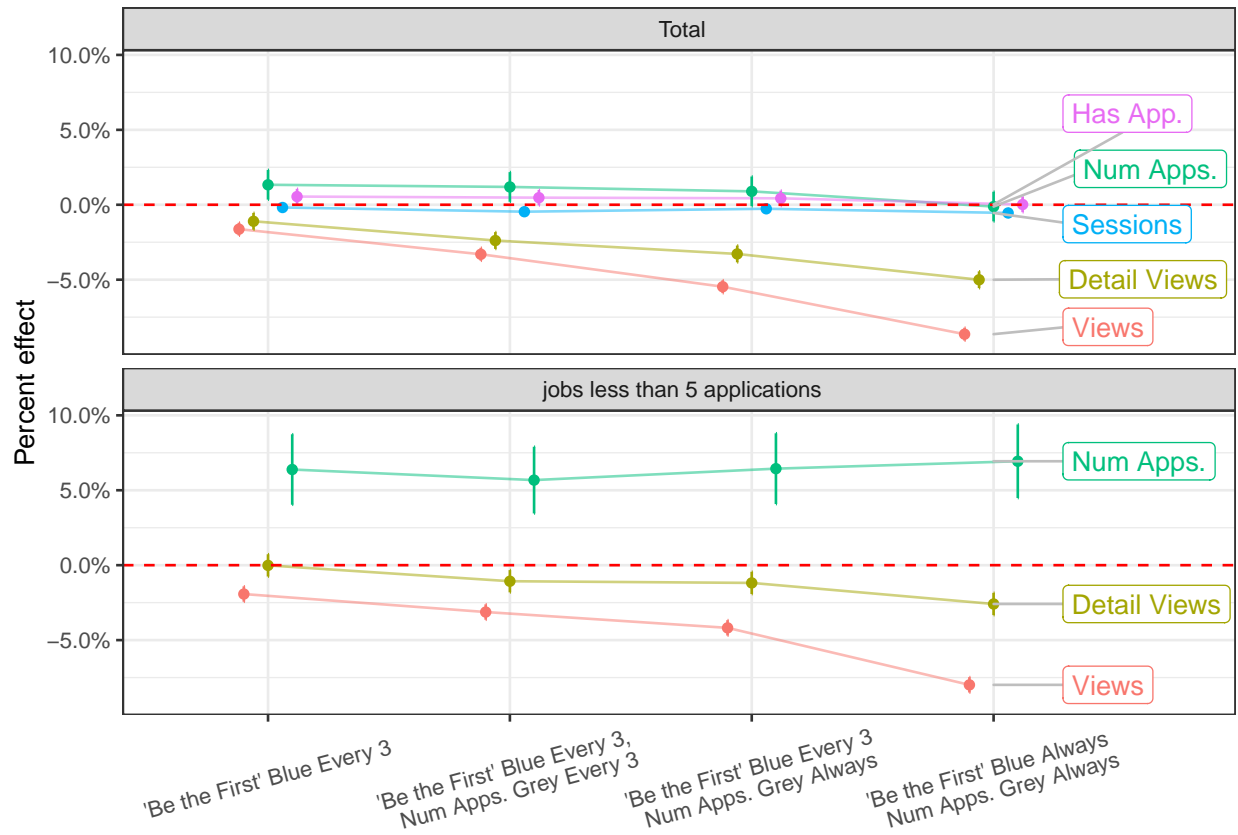


Table B.2: Treatment effects pooled across all three experiments - by application type

	0 - 4 App. (1)	5 - 20 App. (2)	21 - 50 App. (3)	51 - 200 App. (4)	201+ App. (5)
Treatment	0.0023*** (0.0003)	0.0003 (0.0003)	$-9.58 \times 10^{-5}$ (0.0002)	-0.0007* (0.0003)	-0.0001 (0.0002)
Mean of Y:	0.065	0.089	0.069	0.089	0.021
R <sup>2</sup>	0.001	0.001	0.002	0.002	0.010
Observations	29,375,533	29,375,533	29,375,533	29,375,533	29,375,533
Experiment fixed effects	✓	✓	✓	✓	✓

Notes: Effects of the competition signal treatment pooled across experiment.

## Appendix C: Why does the effect size vary across experiments?

We now investigate why the effects of competition information on applications vary so greatly across the three experiments. We show that the details of the treatment implementation, changes in the demographics of users, and changes in market tightness do not explain the differences in treatment effects.

### C.1. Differences in treatment

As explained in Section B, each of our three experiments had several treatment variations. One concern is that our main results are driven by differences in the exact implementation of the treatment across experiments. In this section, we compare two *identical* treatment arms across experiments 2 and 3 and show that the differences in experimental treatment effects persist even for identical treatments.

The first repeated treatment is one in which the ‘Be one of the first to apply’ signal is eligible to be shown in blue every three tiles. The estimates and 95% confidence intervals for this treatment are shown in Comparison A of Figure C.1. The effect of the treatment on applications to under-subscribed vacancies is more than twice as large in experiment 3 than it is in experiment 2 (pval: 0.016). There are also substantial differences in other outcomes, such as the number of detail views and views for all vacancies.

Similarly, there are differences in the effects of the other repeated treatment between experiment 2 and 3. This treatment displayed competition information every 3rd tile for all types of information. Furthermore, ‘Be one of the first to apply’ is shown in blue. The effect of the treatment on applications to under-subscribed vacancies is more than twice as large in experiment 3 than it is in experiment 2 (pval: 0.07). There are also substantial differences in other outcomes, such as the number of detail views and views for all vacancies. As a result, we conclude that the differences in experiments are not driven by the specific implementation of the competition signal.

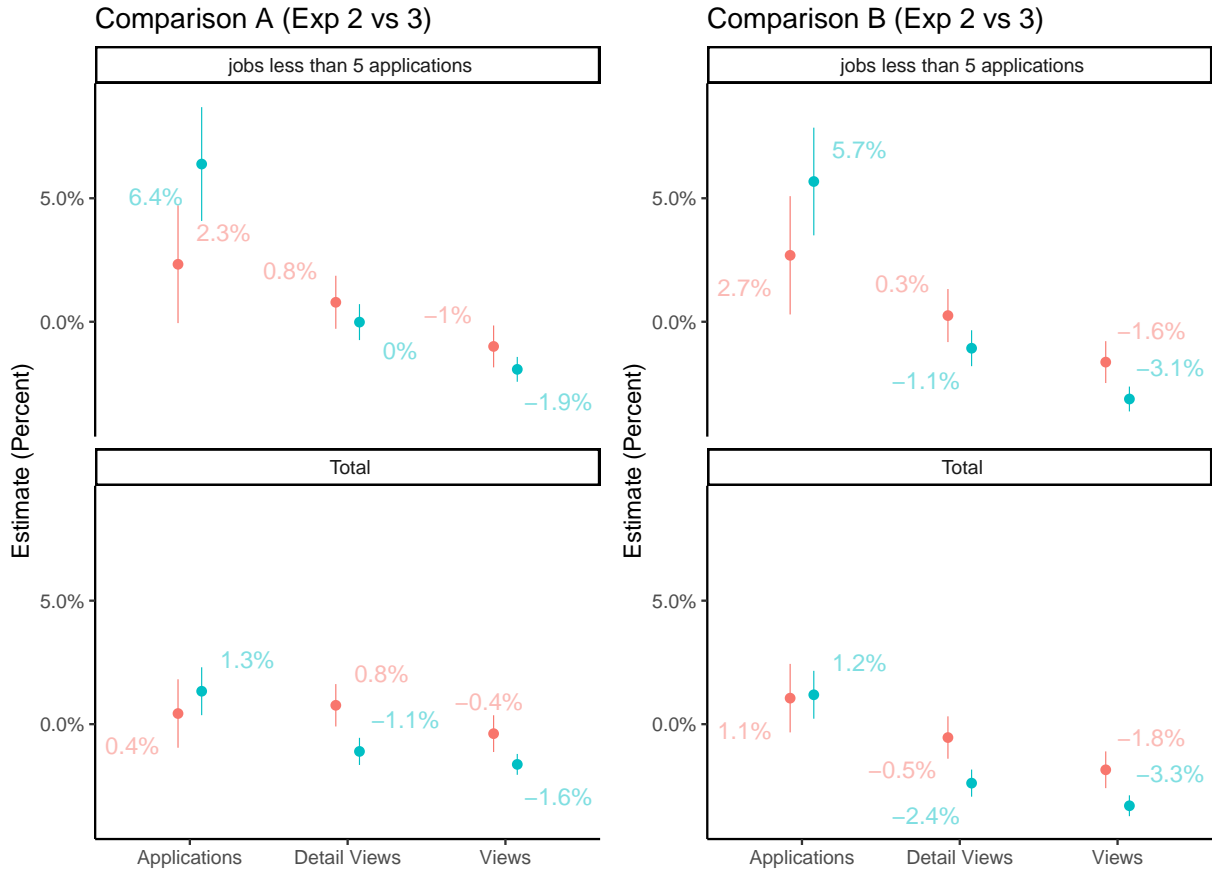
### C.2. Differences in observable user characteristics and market conditions.

Another reason for the differences in treatment effects across experiments may be that the user composition or market conditions are changing. JOF is a fast growing and global platform, so it is conceivable that these factors could change over a period as short as a month.

Table C.1 reports summary statistics for user characteristics for the three experiments. There are some compositional differences across the experiments—for example, by Experiment III, the fraction of users who are from the US has declined, as has the fraction that are female. Furthermore, Experiment III has a lower share of users who had used the Jobs product in the two weeks prior to the experiment than Experiment II. We can also measure the market tightness of each commuting zone in our sample - defined by the prior week’s number of applications divided by number of vacancies. Figure C.2 plots the evolution of this quantity over time and by region. We see that tightness increases after Experiment I and falls after Experiment II.

Next, we test for heterogeneous effects based on these factors, and find that they are not large enough to explain the differences between experiments. We estimate separate regressions interacting a dichotomized version of each variable with the treatment, where the outcome variable is applications to under-subscribed jobs. The results of these regressions are reported in Figure C.3. We see that there is some heterogeneity in treatment effects for those who’ve used the product before and for US users. However, this heterogeneity is not precisely estimated.

Figure C.1: Effects of the same treatment across experiments



### C.3. Differences in viewed vacancies across experiments

We now consider whether differences in the vacancies shown to job seekers can explain the differential effects of the experiments. As in the proceeding section, we first measure whether there is any shift in vacancy characteristics over time. We then test whether there are heterogeneous treatment effects as a result of these characteristics.

Figure C.4 displays the average values of our measured variables across the three experiments (colors) and display position in the interface (x-axis). Our first measure of job characteristics is whether the candidate can apply to the job through Facebook. We see that for all experiments, the share of exposures to these jobs is above 85%. Our second measure is whether the job was created through Facebook’s portal (native) or whether it was syndicated from a third-party. We see that most vacancies are native and that there are fewer third-party vacancies in experiment 1. Our next two measures incorporate information that is displayed on the platform. Namely, we can observe the prior number of applications and vacancy age at the time of viewing. We see some differences, but these are not monotonic in experiment launch date. In particular, Experiment II is less likely to have under-subscribed vacancies be viewed, while Experiment I has a lower vacancy age exposed.

Figure C.2: Evolution of market tightness over time

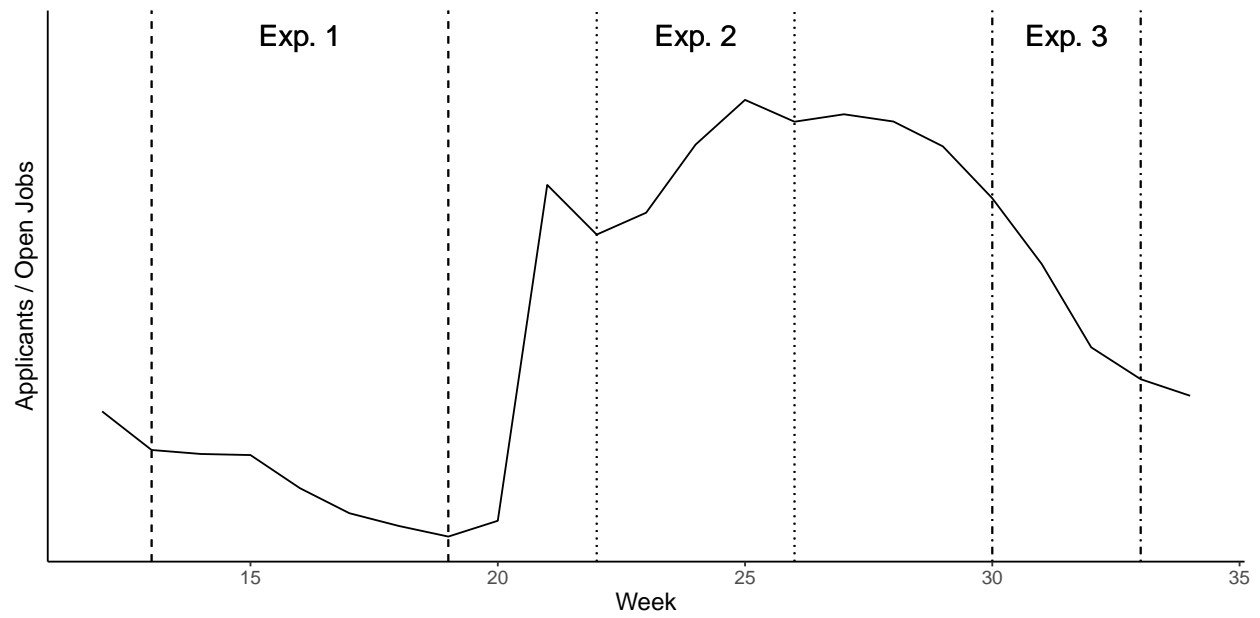


Figure C.3: Heterogeneous treatment effects - Applications to under-subscribed vacancies

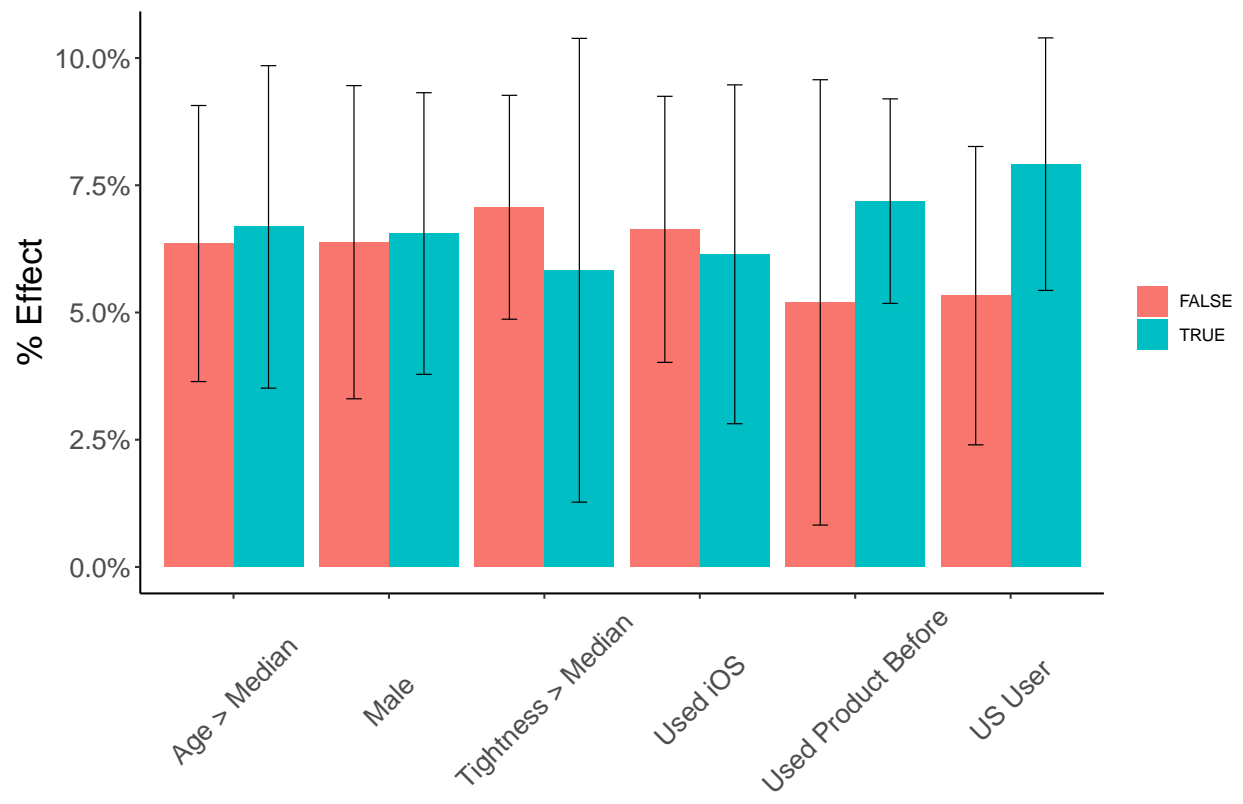


Table C.1: Control group demographics over the three experiments

	25th	Median	75th	Mean	StDEv	exp
Exp I (n = 1,763,735)						
Age	26	33	44	35.95	13.63	1
US User	NA	NA	NA	0.25	NA	1
Friends	219	444	873	746.74	896.28	1
Used Jobs Pre Exp.	NA	NA	NA	0.27	NA	1
iOS User	NA	NA	NA	0.34	NA	1
Male	NA	NA	NA	0.59	NA	1
Exp II (n = 863,215)						
Age	25	32	42	34.45	13.18	2
US User	NA	NA	NA	0.21	NA	2
Friends	203	437	922	785.59	969.06	2
Used Jobs Pre Exp.	NA	NA	NA	0.38	NA	2
iOS User	NA	NA	NA	0.27	NA	2
Male	NA	NA	NA	0.54	NA	2
Exp III (n = 3,265,172)						
Age	24	31	42	34.23	13.74	3
US User	NA	NA	NA	0.20	NA	3
Friends	173	399	873	747.61	962.37	3
Used Jobs Pre Exp.	NA	NA	NA	0.32	NA	3
iOS User	NA	NA	NA	0.31	NA	3
Male	NA	NA	NA	0.51	NA	3

*Notes:* User characteristics by experiment. ‘Used Jobs Pre Exp.’ is an indicator for whether the user used the job board in the two weeks prior to the experiment.

Next, we consider whether these characteristics can explain treatment effect heterogeneity in Experiment III. [Table C.4](#) displays regressions where the outcome is whether the user applied to the first vacancy viewed in the experiment. We exclude vacancies shown in positions below the first because the depth of search is an endogenous outcome and we exclude vacancies that could not be applied to through Facebook.<sup>19</sup>

The explanatory variables are vacancy and vacancy by user characteristics. Column (1) displays the results while including the variables considered in [Figure C.4](#). We see that these variables do not explain treatment effect heterogeneity. Column (2) further adds the characteristics of the vacancy poster - in particular whether the poster had a ‘local’ page on Facebook and whether they had a ‘business’ page on Facebook. It also adds an indicator for whether the vacancy and the searcher are located in the same city. The interactions are not statistically significant for any of these variables.

Finally, in column (3), we add the algorithmic score which predicts whether a given searcher will apply to a particular job. This score is generated using a proprietary machine learning algorithm and does substantially

<sup>19</sup> This designation sometimes changed over time and a negligible fraction of these jobs did have an application.

Table C.3: Differences conditional on application order

	Viewed	Contact	Interview
Treatment	0.001	0.002*	0.000
	(0.001)	(0.001)	(0.000)
Log(Order)	-0.018***	-0.004***	-0.003***
	(0.000)	(0.000)	(0.000)
Treat * Log(Order)	0.000	-0.001**	0.000
	(0.000)	(0.000)	(0.000)
N	5472586	5472586	5472586
Vacancy FE	X	X	X

*Notes:* This table contains results for a linear regression of applications outcomes on the treatment in experiment 3. ‘Viewed’ is an indicator whether the employer viewed the application, ‘Contact’ is an indicator for whether an employer sent an applicant a message, and ‘Interview’ is an indicator for whether an employer marked that an interview was conducted.

Figure C.4: Characteristics of vacancies viewed across experiments

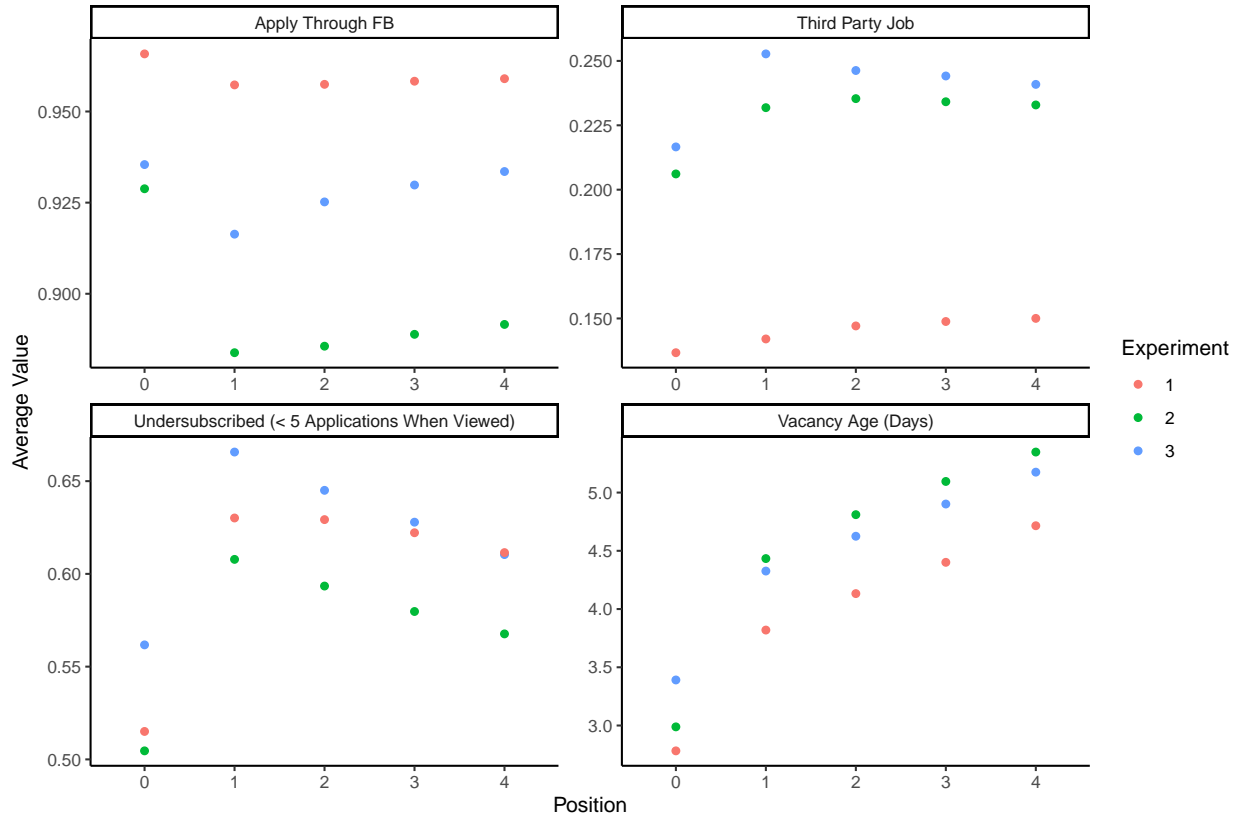


Table C.4: Characteristics of vacancies viewed across experiments

	Has Application		
	(1)	(2)	(3)
Treatment	0.0004*** (1e-04)	-0.0002 (4e-04)	-0.0006 (0.0004)
Treatment * Third Party	0.0000 (1e-04)	0.0001 (1e-04)	0.0001 (0.0001)
Treatment * Age > 5 Days	0.0000 (1e-04)	0.0000 (1e-04)	0.0002 (0.0001)
Treatment * Same City		0.0001 (1e-04)	0.0001 (0.0001)
Treatment * FB Local		0.0000 (1e-04)	0.0001 (0.0001)
Treatment * FB Business		0.0005 (4e-04)	0.0006* (0.0004)
Treatment * Algorithmic Match Score			0.0053*** (0.0012)
Num.Obs.	9020389	9020389	9020389
R2	0.000	0.000	0.004
R2 Adj.	0.000	0.000	0.004

*Notes:* Regression of whether the user applied to the first vacancy shown as a function of the treatment and interactions. Non-interacted covariates are not shown. Observations for which the first job shown could not be applied to through Facebook were excluded.

predict treatment effect heterogeneity. Furthermore, this effect is an order of magnitude larger than the average treatment effect. This heterogeneity demonstrates that there is scope for the match between searcher and vacancy to change over time and to cause treatment effect heterogeneity.



## Appendix D: Survey Choice Experiment

The pre-registered survey choice experiment consists of the following questions.<sup>20</sup> The first module asks about the employment status, age, and gender of a respondent, whether the respondent is actively looking forward, and an attention check. The survey then consists of three comparisons with two jobs each. Figure D.1 displays the three choice scenarios for one realization of the random draws. Each choice is between two companies, Blank Co and Brown Co, which differ in their wages, number of current applications, and an AI probability that the respondent gets an offer.

There are four elements of the survey that are randomized. First, some participants see information about an artificial intelligence (AI) probability that they receive an offer for a job while others do not. Whether this information is shown is randomized at the respondent level. There are three additional randomizations, one for each choice scenario. In each choice scenario, whether Blank Co or Brown Co has the lower number of applications is randomized at a question by respondent level.

There are then several post-choice scenario questions. For each choice in which a respondent answers that they either prefer Blank Co or Brown Co, the respondent is asked to explain their choice in a text box. Note that no such question is asked when the response is ‘No Preference’. After the open text responses, we finish the survey by asking whether the participant responded randomly and whether the participant has feedback about the survey.

We now describe additional analysis details that we mentioned in our pre-registration. The experimental sample contained 1189 respondents, of which 592 were in the condition without an AI probability displayed. We investigate the experiment through regression analysis, displayed in Table D.1. The outcome in all of the specifications is a variable that is coded as -1 if Blank Co was chosen, 0 if there was no preference, and 1 if Brown Co was chosen. Column 1 displays the baseline regression with standard errors clustered at the participant level and shows that Brown Co is chosen more often when it has fewer applications.

Next, we consider the effect of information about the AI predicted probability of offer. Column 2 displays results with an intercalation between the main treatment (lower applications) with whether the AI probability was shown. The coefficient on the interaction is negative, demonstrating that information about prior applications has less of an effect when the probability of an offer is known. However, there is still some effect of the information even in the AI condition. We can reject the null of no effect in the AI group with a Wald Test ( $p < 3.4e-31$ ).

Lastly, we consider heterogenous treatment effects. Column 3 displays the effect of the treatment separately for each comparison. We find that for each question, respondents prefer vacancies with fewer prior applications. Columns 4 and 5 estimate heterogeneous effects by gender and whether the respondent searched for a job in the past year. We find that there is no statistically significant differences in responses by gender, but that there are differences by whether the respondent searched for a job. In particular, those who searched for a job have a stronger preference for vacancies with fewer prior applicants than those who did not search for a job in the past year.

<sup>20</sup> The experiment was determined to be exempt from the IRB by MIT’s Committee on the Use of Humans as Experimental Subjects. The preregistration for the experiment is available here: <https://www.socialscienceregistry.org/trials/9344>.

Figure D.1: Survey choice questions

**BOSTON UNIVERSITY**

Imagine that you are currently unemployed and are searching for a job:

- Each job requires you to work for 40 hours a week.
- Each job was posted at the same time.
- Each job is identical except for the information displayed to you.
- You would take any of the jobs if it was the only offer you got.

For each job, please state whether you would value the ability to apply for Blank Co or Brown Co more. Our job search engine uses an artificial intelligence (AI) algorithm that uses all information to predict whether you would receive an offer if you applied. This probability is also included in the job description.

**Blank Co**  
\$18/hr  
Currently has: 200 + Applications  
AI probability you get an offer: 25%

**Brown Co**  
\$20/hr  
Currently has: 5 - 20 Applications  
AI probability you get an offer: 25%

Blank Co  
☐

No Preference  
☐

Brown Co  
☐

**BOSTON UNIVERSITY**

Imagine that you are currently unemployed and are searching for a job:

- Each job requires you to work for 40 hours a week.
- Each job was posted at the same time.
- Each job is identical except for the information displayed to you.
- You would take any of the jobs if it was the only offer you got.

For each job, please state whether you would value the ability to apply for Blank Co or Brown Co more. Our job search engine uses an artificial intelligence (AI) algorithm that uses all information to predict whether you would receive an offer if you applied. This probability is also included in the job description.

**Blank Co**  
\$19/hr  
Currently has: 0 - 4 Applications  
AI probability you get an offer: 13%

**Brown Co**  
\$21/hr  
Currently has: 200 + Applications  
AI probability you get an offer: 13%

Blank Co  
☐

No Preference  
☐

Brown Co  
☐

(a) Comparison Question 1

(b) Comparison Question 2

**BOSTON UNIVERSITY**

Imagine that you are currently unemployed and are searching for a job:

- Each job requires you to work for 40 hours a week.
- Each job was posted at the same time.
- Each job is identical except for the information displayed to you.
- You would take any of the jobs if it was the only offer you got.

For each job, please state whether you would value the ability to apply for Blank Co or Brown Co more. Our job search engine uses an artificial intelligence (AI) algorithm that uses all information to predict whether you would receive an offer if you applied. This probability is also included in the job description.

**Blank Co**  
\$22/hr  
Currently has: 0 - 4 Applications  
AI probability you get an offer: 30%

**Brown Co**  
\$24/hr  
Currently has: 5 - 20 Applications  
AI probability you get an offer: 30%

Blank Co  
☐

No Preference  
☐

Brown Co  
☐

(c) Comparison Question 3

*Notes:* Survey choice questions. Note that whether the higher application count was displayed for Blank Co or Brown Co was randomized at a question by respondent level. Whether the line about the AI probability was shown was randomized at a respondent level.

Table D.1: Survey Regressions

	Choice (-1 (Blank), 0, 1 (Brown))				
	(1)	(2)	(3)	(4)	(5)
(Intercept)	-0.0238 (0.0364)	-0.0238 (0.0363)		-0.0136 (0.0501)	0.0357 (0.0576)
Fewer Applications (Brown)	0.8704*** (0.0422)	0.8704*** (0.0422)		0.8469*** (0.0587)	0.7494*** (0.0692)
AI Probability		0.4051*** (0.0517)			
Fewer Applications (Brown) $\times$ AI Probability		-0.3736*** (0.0583)			
Fewer Applications (Brown) $\times$ Question = 1			0.9806*** (0.0585)		
Fewer Applications (Brown) $\times$ Question = 2			1.140*** (0.0610)		
Fewer Applications (Brown) $\times$ Question = 3			0.4925*** (0.0558)		
Male				-0.0203 (0.0727)	
Fewer Applications (Brown) $\times$ Male				0.0450 (0.0844)	
Searched for Job					-0.1012 (0.0742)
Fewer Applications (Brown) $\times$ Searched for Job					0.2076** (0.0867)
Observations	1,776	3,567	1,776	1,776	1,776
R <sup>2</sup>	0.25896	0.21245	0.31535	0.25913	0.26254
Within R <sup>2</sup>			0.29417		
Sample	No AI	All	No AI	No AI	No AI
Question fixed effects			✓		

Notes: The outcome for all regressions is a variable that is coded as -1 if Blank Co was chosen, 0 if there was no preference, and 1 if Brown Co was chosen

Figure D.2: Distribution of responses  
set of choices with AI probability displayed

